# ***TPMT*, *COMT* and *ACYP2* genetic variants in paediatric cancer patients with cisplatin-induced ototoxicity**

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# **Abstract**

**Objectives**

Cisplatin ototoxicity affects 42-88% of treated children. *COMT*, *TPMT*, and *AYCP2* genetic variants have been associated with ototoxicity, but the findings have been contradictory. The aims of the study were to (a) investigate these associations in a carefully phenotyped cohort of UK children, and (b) to undertake a systematic review and meta-analysis.

**Methods**

### We recruited 149 children from seven UK centres using a retrospective cohort study design. All participants were carefully clinically phenotyped. Genotyping was undertaken for one *ACYP2* (rs1872328), three *TPMT* (rs12201199, rs1142345 and rs1800460), two *COMT* (rs4646316 and rs9332377) variants.

**Results**

For CTCAE grading, hearing loss was present in 91/120 (75.8%; worst ear) and 79/120 (65.8%; better ear). Using Chang grading, hearing loss was diagnosed in 85/119 (71.4 %; worst ear) vs 75/119 (63.0%; better ear). No *TPMT* or *COMT* SNPs were associated with ototoxicity. *ACYP2* SNP rs1872328 was associated with ototoxicity (p=0.027; worst ear). Meta-analysis of our data with that reported in previous studies showed the pooled odds ratio (OR) to be statistically significant for both the *COMT* SNP rs4646316 (odds ratio 1.50, 95% CI: 1.15-1.95) and the *ACYP2* SNP rs1872328 (odds ratio 5.91, 95% CI: 1.51-23.16).

**Conclusions**

In conclusion, we showed an association between the *ACYP2* polymorphism and cisplatin-induced ototoxicity, but not with the *TPMT* and *COMT*. A meta-analysis was statistically significant for both the *COMT* rs4646316 and the *ACYP2* rs1872328 SNPs. Grading the hearing of children with asymmetric hearing loss requires additional clarification.

**Keywords**

Cisplatin, Paediatric cancer, Cancer, Pharmacogenetics, Ototoxicity, *ACYP2*, *TPMT*, *COMT*

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# **Introduction**

Cisplatin is a chemotherapeutic agent used to treat solid malignancies in childhood, including in the central nervous system (CNS). Although highly effective, its therapeutic index is narrow. Adverse effects include irreversible bilateral sensorineural hearing loss [1].

Hearing loss in childhood affects crucial areas of development. Even minimal sensorineural hearing loss can affect both academic ability and overall level of function (behaviour, energy, stress, social support, and self-esteem) [2]. It may be of particular importance for children with brain tumours where a variety of neuro-toxicities can occur and significantly combine to impair recovery and lead to worsening disability.

There are established risk factors for the development of cisplatin related hearing loss including cumulative dose of cisplatin, younger age (especially <5 years), concomitant (or preceding) radiotherapy to the CNS [3], and exposure to carboplatin in myelo-ablative doses [1,4-6].

The reported incidence rates for cisplatin induced hearing loss in children range between 42-88% [1,5,7-18]. Study populations have often been small (median 67 patients; range 22-238), and heterogeneous with regard to diagnoses, age, dose of cisplatin, treatment schedules, hearing grading and co-administration of concurrent ototoxic agents and cranial radiotherapy. In addition, there is no consensus about how to define hearing loss, leading to variability in the assessment and grading of ototoxicity.

Despite this, there does also appear to be significant inter-individual variability in cisplatin induced hearing loss. Irreversible hearing loss can occur after a single dose of cisplatin in some individuals, whereas other children do not develop hearing loss even after multiple and high doses of cisplatin [9]. Several studies have identified potential predisposing genetic variants influencing cisplatin induced hearing loss. Thiopurine Methyltransferase (*TPMT*) and Catechol-O-Methyltransferase (*COMT*) [12] (n=162) risk genotypes were first described in 2009, with *TPMT* (but not *COMT*) being replicated in a subsequent cohort [15] (n=155). However neither gene was replicated in three additional studies [16,17] (n=213, n=110, and n=38). A meta-analysis of all of the above studies did associate *COMT* (but not *TPMT*) with cisplatin induced ototoxicity [17]. Subsequently, another group have failed to find an association with either *TPMT* or *COMT* [18] (n=63). There is therefore considerable uncertainty about whether *TPMT* and/or *COMT* polymorphisms represent genuine risk factors for cisplatin-induced ototoxicity.

More recently, another genetic variant (*ACYP2* (rs1872328)) was identified using a genome-wide association study in 238 children with brain tumours [14], which was replicated within the original study. An additional study has also recently replicated the association with *ACYP2* in patients (n=156) with osteosarcoma [19]. In the present study, our aim was two-fold: first, to test for an association between variants in the *COMT*, *TPMT* and *ACYP2* genes and cisplatin induced hearing loss, in a carefully phenotyped cohort of UK children; and second, to undertake a systematic review and meta-analysis to determine the association with *TPMT*, *COMT* and *ACYP2* data.

## **Materials and Methods**

### **Study design and criteria**

Participants were recruited to the Molecular Genetics of Adverse Drug Reactions in Childhood (MAGIC) study. Ethical approval for this study was granted by North West 3 Research Ethics Committee (10/H1002/57). For cisplatin, a target recruitment of 400 has been established for a genome wide association study. This sub-study sample size was determined by recruitment achieved at the time the candidate gene analysis was undertaken.

For inclusion into the study, the patients needed to have started cisplatin on or after 1st January 2001 and had at least one evaluable audiogram following the last dose of cisplatin (post treatment audiogram). To be considered evaluable the audiogram had to fulfil the following criteria: (1) Either pure tone audiogram (PTA) or visual response audiogram (VRA) in db HL and (2) tested at 1, 2 and 4 kHz and either 6 or 8 kHz. The exclusion criteria were as follows:

1. Parent/guardian unwilling to take part (if participant <16 years at recruitment).
2. Participant unwilling to consent (if ≥16 years at recruitment).
3. Competent paediatric participant unwilling to assent (competence assessed on a case by case basis).
4. Hearing impairment prior to cisplatin treatment.
5. No evaluable post treatment audiogram
6. Patient is, in the opinion of the investigator or the clinical team, not suitable to participate in the study.

### **Consent and data collection**

All parents (patient <16 years) or patients (age ≥16 years) provided written informed consent prior to recruitment to this study. Hearing impairment was assessed on the post-treatment audiogram.

### **DNA collection and extraction**

Patient samples for DNA were collected as whole blood ethylenediaminetetraacetic acid (EDTA) samples or salivary samples. DNA collection and extraction for saliva samples has been described previously [20]. EDTA blood samples were stored at -80°C and following defrosting, genomic DNA was extracted using the Chemagen whole-blood DNA extraction kit on the Chemagic Magnetic Separation Module I according to the manufacturer's protocol (PerkinElmer chemagen Technologie GmbH, Baesweiler, Germany; www.chemagen.com).

### **Genotyping**

Genotyping was undertaken for three *TPMT* variants (rs12201199, rs1142345 and rs1800460) and two *COMT* variants (rs4646316 and rs9332377) [12], and the SNP rs1872328 for *ACYP2* [14], as described previously [20]. The following Taqman Drug metabolism genotyping assays were used: C\_\_31923406\_10 (for rs12201199), C\_\_19567\_20 (for rs1142345), , C\_\_30634116\_20 (for rs1088460), C\_\_29193982\_10 (for rs4646316), C\_\_29614343\_10 (for rs9332377) and C\_\_11643398\_10 (for rs1872328) .

### **Phenotype Definition**

We graded all audiograms according to CTCAE for *COMT* and *TPMT* [12,21] and assigned Chang grades for *ACYP2* [14,22] (Supplementary Table S1). If several post treatment audiograms were available, the highest-grade audiogram was used in the analysis. For patients with asymmetric hearing loss, we graded both ears separately and the results of both grades analysed as outlined below.

### **Statistical Analysis**

Quality control procedures were applied to the genotype data and individuals or SNPs included in the analysis based on the following criteria: sample call rate (samples missing ≥2 SNPs excluded), SNP call rate (only SNPs with call rate >95% included), minor allele frequency (only SNPs with minor allele frequency (MAF) > 0.01 included) and Hardy-Weinberg (HW) test (only SNPs with Hardy-Weinberg test p-value > 0.05 included). An additive mode of inheritance was assumed with SNPs coded 0, 1 or 2 to represent wild-type homozygotes, heterozygotes and mutant-homozygotes respectively.

For the purpose of our primary analyses, our phenotype was treated as ordinal. First, a univariate multinomial logistic regression model was fitted for each non-genetic factor in turn, to identify which non-genetic factors to adjust for in the SNP association analyses. Next, multivariable multinomial logistic regression models were fitted for each SNP in turn. For each SNP, two models were fitted. The first model included covariates to represent all non-genetic factors with p<0.25 univariately. Stepwise variable selection was applied to this baseline model to remove any covariates no longer significant in the multivariable model. The final model following variable selection was called the ‘baseline model’. The second model (‘the genetic model’) was the same as the baseline model but also included a covariate to represent the SNP. The likelihood ratio test was applied to compare the two models and thus assess for statistical significance of the SNP. Since 5 SNPs were tested for association with the outcome of CTCAE grade, a Bonferroni adjustment for 5 tests was applied to these analyses of association. No adjustment was applied to the test for association between the *ACYP2* SNP and Chang grade. In cases of asymmetric hearing loss, the worse ear grade was used as final ototoxicity grade [14,16,22]. In order to avoid bias arising from this approach, a sensitivity analysis was carried out using the ototoxicity grade of the better ear as final grade. Further sensitivity analyses of ototoxicity grades were performed by dichotomising outcomes in three different ways: CTCAE grade 0 vs. 1-4; CTCAE grade 0 vs. 2-4; CTCAE grade 0 vs. 3-4. The statistical approach to the sensitivity analyses was the same as for the ordinal outcome but logistic regression models were used instead of multinomial logistic regression models.

**Systematic Review and Meta-Analysis**

A search was undertaken on 17th March 2016 in EMBASE and MEDLINE databases using the search strategy and exclusion criteria detailed in Supplementary Table S9. The paediatric search terms were based on published examples [23].

Two reviewers (CB and ALJ) independently screened all papers identified for inclusion. Any conflicts were resolved by discussion (CB and ALJ). Methodological quality of the papers was assessed using a published quality assessment checklist [24]. The following data were extracted from each study: year of publication, ethnicity of participants, SNPs investigated, outcomes investigated including their definition, sample size, and study design. If >1 study investigated associations between the same SNP and outcome combination, data required to undertake a meta-analysis was extracted, including: numbers in each genotype group, number of cases and controls per genotype group, odds ratio, standard error of odds ratio, confidence interval for odds ratio.

These data were synthesised using the software package Review Manager 5.3 (The Cochrane Collaboration, 2014). Since different papers had undertaken analyses assuming different modes of inheritance, and due to the variability between studies in how data were reported, it was only possible to conduct a meta-analysis where the allelic odds ratio was calculated (i.e. the odds ratio of developing ototoxicity for the mutant allele vs wild-type allele). The statistical method used to estimate a pooled odds ratio across studies was the Mantel-Haenszel random-effects method [25] and heterogeneity was assessed by referring to the I2 statistic [26,27]. No formal adjustment for study quality was made in the meta-analyses; however results of assessing methodological quality were considered when exploring potential sources of heterogeneity.

# **Results**

One hundred and forty-nine patients were consented, but data was not available for six. In addition, 23 did not have the required audiograms to allow any grading of ototoxicity, and a single patient’s data were insufficient to distinguish between Chang grade 1b and 2a. There were 120 evaluable patients for CTCAE and 119 patients for the Chang criteria (Supplementary data Figure S2).

The distribution of underlying diagnoses was: Medulloblastoma (30.0%; 36/120), Hepatoblastoma (12.5% 15/120), Osteosarcoma (24.2%; 29/120), Neuroblastoma (12.5%; 15/120), other CNS tumours (15.8%; 19/120), and other non-CNS tumours (5.0% 6/120). The self-reported ethnicity for 88.3% (106/120) patients was Caucasian, for 5% Asian (6/120), for 3.3% (4/120) African, for one child Chinese and for three children, the ethnicity was not known.

Clinically, using CTCAE grading, 91/120 patients (75.8%) experienced hearing loss (analysis of worst ear) vs 79/120 patients (65.8%) (better ear). Using Chang grading and considering the worse ear in cases of asymmetrical hearing loss, 85/119 patients (71.4 %) experienced hearing loss vs 75/119 (63.0%) considering the better ear. In our study, the number of patients with asymmetric hearing loss, leading to differential grading in each ear was therefore 12/120 (10.0 %) for CTCAE grading and 10/119 (8.4%) for Chang grading.

### **Genomic Quality control results**

Four patients were removed from the analysis after quality control as results for two or more SNPs were missing. All variants had minor allele frequency >5% and all passed the HW test (p values > 0.05). Table 1 shows the demographic details of the 116 children included in the genetic analysis.

### **Results of univariate analysis**

Results of univariate analyses of association with each non-genetic factor are also shown in Table 1. Of the 98 children in this study who did not receive combined cisplatin and carboplatin therapy, 27 received carboplatin after cisplatin therapy, likely as alternative therapy due to cisplatin induced nephro- or ototoxicity. All children in this group also had grade CTCAE 1-4 hearing loss (worse ear), whereas 1/27 did not have hearing loss using Chang criteria (worse ear) and two children did not have hearing loss on the better ear (Chang and CTCAE criteria).

### **Analysis of *COMT* and *TPMT* variants**

The summary statistics detailing frequency of hearing loss for each genetic variant investigated are shown in Table 2. Clinical factors included in the multivariable analysis (p<0.25) were: patient age at diagnosis, gender, cranial irradiation, cumulative dose of cisplatin, exposure to vincristine and carboplatin (Table 1). On applying variable selection to the multivariable model including all these factors, vincristine was removed due to correlation with cranial irradiation (correlation = 0.52). The Bonferroni corrected likelihood ratio test (LRT) p-values for the 5 SNPs are shown in Table 3. None of the *TPMT* or *COMT* SNPs were associated with cisplatin induced hearing loss (Table 3). In sensitivity analyses, there were still no significant associations with cisplatin induced hearing loss for any of the variants using dichotomised outcomes (Supplementary tables S3+S4), better ear grade for ordinal outcomes (Supplementary table S5) or when restricting the analysis to individuals of Caucasian ancestry (Supplementary tables S6,S7and S8).

### **Analysis of *ACYP2* variant**

The summary statistics detailing frequency of hearing loss for each genetic variant investigated are shown in table 2. Clinical factors included in the multivariable analysis (p<0.25) were: patient age at diagnosis, gender, cranial irradiation, cumulative dose of cisplatin, exposure to vincristine and carboplatin (Table 1). *ACYP2* SNP rs1872328 was associated with ototoxicity (p=0.027), where ototoxicity grade was modelled as an ordinal variable with reference to the worse ear grade in cases of asymmetric hearing loss (Table 4). However, in sensitivity analyses using the better ear grade and ordinal as well as dichotomised outcomes, there was no significant association between the investigated *ACYP2* variant and cisplatin induced hearing loss (Supplementary data tables S9, S10, and S11). When restricting the analysis to individuals of Caucasian origin, the association was no longer significant (Table 5).

### **Systematic Review and Meta-Analysis**

The search strategy identified 256 possibly relevant papers, but after screening for inclusion/exclusion criteria 7 studies were included [12,14-19]. The Quorum flow chart detailing the studies identified and included is shown in Supplementary Figure S1, with key study information summarised in Table 6. Results of the methodological quality assessment of included studies are in Supplementary Table S10, with discussion of these results in Supplementary Appendix S11.

It was possible to undertake six different meta-analyses, one each for SNPs rs12201199, rs1800460 and rs1142345 in the *TMPT* gene, one each for SNPs rs4646316 and rs9332377 in the *COMT* gene and one for SNP rs1872328 in the *ACYP2* gene. For the first five meta-analyses, cases were defined as those with CTCAE grade 2-4 whilst controls were defined as those with CTCAE grade 0. For the sixth meta-analysis, cases were defined as those with Chang grade >0 whilst controls were defined as those with Chang grade 0. Although the study by Lanvers-Kaminsky, investigating the rs12201199 and rs9332377 SNPs, used a different method for defining ototoxicity (Muenster classification), it was included in the meta-analyses. However, sensitivity meta-analyses excluding the study by Lanvers-Kaminsky were also conducted for these two SNPs. Forest plots illustrating the results of the meta-analyses are provided in Figures 1 A-F.

The pooled odds ratio was statistically significant for the associations with the *COMT* SNP rs4646316 (odds ratio 1.50, 95% CI: 1.15-1.95) and with the *ACYP2* SNP rs1872328 (odds ratio 5.91, 95% CI: 1.51-23.16). In both cases, the extent of heterogeneity between studies was low (3% and 6% respectively). For the rs12201199 SNP in the *TPMT* gene, the pooled odds ratio was not statistically significant (odds ratio 1.83, 95% CI: 0.89-3.76) but there was relatively high heterogeneity between studies (I2: 59%). The same was true for the other two *TPMT* SNPs, rs1142345 (odds ratio 1.59, 95% CI: 0.67-3.77, I2: 53%) and rs1800460 (odds ratio 1.34, 95% CI: 0.54-3.35, I2: 47%). The pooled odds ratio for *COMT* SNP rs9332377 was also non-significant (odds ratio 1.15, 95% CI: 0.67-1.95) but again the level of heterogeneity between studies was relatively high (I2:66%). When excluding the study by Lanvers-Kaminsky from the meta-analyses for SNPs rs12201199 and rs9332377 the conclusions remained the same (data not shown), but the level of heterogeneity increased in both cases.

# **Discussion**

Our study highlights known risk factors for cisplatin induced ototoxicity in children, including increasing cumulative dose of cisplatin [1,5,9,11,28], younger age [5,11,16], cranial radiotherapy [4,13,29] and exposure to other ototoxic agents such as vincristine [15]. Our study did not detect ethnic origin as a risk factor for hearing loss in patients receiving cisplatin therapy, but this was not our intention as nearly 90% of our study population were Caucasian.

Our data did not replicate previous findings that *COMT* and *TPMT* variants are risk factors for cisplatin-induced ototoxicity, consistent with several other studies [16-18]. Heterogeneity between study populations with regards to size, age range, tumour type, cranial radiotherapy, ethnicity, use of potentially otoprotectant therapy (amifostine) and retrospective versus prospective nature of studies are likely to be confounding variables [16,17]. We addressed these through additional sensitivity analyses, and by using multivariable models, but the conclusions remained the same. Sample size is perhaps the most limiting factor of our study. However, combining our data with previous studies, the pooled odds ratio was statistically significant for the associations with the *COMT* SNP rs464316 (odds ratio 1.50, 95% CI:1.15-1.95, I2:3%), which supports the findings from a previous meta-analysis [17].

In addition, our study replicated the association between the *ACYP2* polymorphism and cisplatin induced ototoxicity, first identified through a GWAS [14] using ordinal outcome measures (Table 4). However, the association was lost when restricting our analysis to individuals of Caucasian ancestry. This is likely due to the “loss” of three cases of ototoxicity in GA carriers (data not shown). We were also not able to demonstrate the association with dichotomised outcomes comparing Chang 0 vs > 0 (Supplementary data section), which was the approach used by Xu et al. [14] to define their discovery cohort. This is likely to be due to several differences between our population and that used by Xu et al [14], including a higher rate of cranial irradiation, higher rate of vincristine use (100% vs 52.2%), a more homogeneous patient group in terms of tumour type, and a lower rate of ototoxicity (73% v 61% Chang grade >0, and 45.2% v 37% Chang ≥2a) in the latter. Furthermore, more than 50% of the patients in the Xu et al cohort received the otoprotectant amifostine compared with none of our patients.

Despite these confounders, we followed the analysis of our primary data by undertaking a meta-analysis comparing Chang 0 vs >0 in rs1872328 ‘AG’ or ‘AA’ genotype carriers vs ‘GG’ carriers as reported in Vos et al [19] as presented in Figure 1. Our pooled odds ratio showed a significant association with this *ACYP2* SNP (odds ratio 5.91, 95% CI: 1.51-23.16).) with no heterogeneity observed (I2: 0%). Taken collectively, our results suggest that there is a true association between *ACYP2* and cisplatin induced ototoxicity, although further studies are still needed to understand how confounding factors affect this association. Furthermore, functional studies should be undertaken to understand the biological plausibility underlying the association, as the explanations to date have been speculative.

Patients in this study who received carboplatin and cisplatin combined, were not at a higher risk of experiencing hearing loss (Table 2). This is not an unexpected result. The majority of children who experience hearing loss after carboplatin therapy have also received cisplatin and/or have received high-dose carboplatin regimens prior to stem cell transplant [6,30]. Children who receive standard dose carboplatin alone experience no, or only mild hearing loss [31]; indeed carboplatin is often used as alternative to cisplatin once significant ototoxicity has been confirmed.

One of the limitations of this study is that the retrospective study design did not make it feasible to collect detailed dosage data of concomitant medications. In common with many of the previous studies, we did not investigate exposure to aminoglycosides, or furosemide in our study. Although ototoxicity is listed in the adverse drug reaction profile in the Summary of Product Characteristics (SPC) for these medicines independent of cisplatin use, studies that did include some or all of them, did not find an association with cisplatin-induced hearing loss [6,12,15,28].

The level of heterogeneity noted in the systematic review between studies examining all of the *TPMT* SNPs, and *COMT* rs9332377 was high (I2 47%-66%). There are several potential sources for this, including different follow up periods between studies, differing study designs, different ethnic groups represented in the study populations, different treatment regimens, and different ages of children recruited. In addition, a factor that has not previously been discussed, but which will impact on all studies investigating ototoxicity, is the grading of asymmetric hearing loss, i.e. worse ear vs. better ear. UK clinical practice has been to use Brock ototoxicity grading, using the better ear to assign the overall grade. Chang, CTCAE and the new SIOP Boston scale [8] do not stipulate which ear to use. Both Xu et al. [14] and Yang et al. [16] used the worse ear to grade ototoxicity but it is not clear in other study populations. This variability in how the hearing loss is recorded may partially account for the difficulties in replicating studies.

In conclusion, we have found an association between the *ACYP2* polymorphism and cisplatin induced ototoxicity, although we could not replicate the association with *TPMT* and *COMT* variants. Cisplatin is used in a wide variety of tumours, and patient heterogeneity is thus likely to be a confounding factor. However, we were able to show in a meta-analysis that there was an association with the *COMT* rs464316 SNP and the *ACYP2* SNP. Further studies in larger populations would still be worthwhile in order to define factors that modulate this association, and importantly, we also need to understand the biological basis of the genetic associations.

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