**Title Page**

Title: The backwards comparability of wrist worn GENEActiv and waist worn ActiGraph accelerometer estimates of sedentary time in children

**Authors:** Lynne M. Boddy1, Robert J. Noonan2, Alex V. Rowlands3,4,5, Liezel Hurter1, Zoe R. Knowles1, Stuart J. Fairclough2.

**Institution and affiliations:**

1Physical Activity Exchange, Research Institute for Sport and Exercise Sciences, Liverpool John Moores University, Liverpool, UK

2Department of Sport and Physical Activity, Edge Hill University, Ormskirk, UK

3Diabetes Research Centre, University of Leicester, Leicester General Hospital, Leicester, UK;

4NIHR Leicester Biomedical Research Centre, UK;

5Alliance for Research in Exercise, Nutrition and Activity (ARENA), Sansom Institute for Health Research, Division of Health Sciences, University of South Australia, Adelaide, Australia.

**Corresponding author:** Dr Lynne M. Boddy, L.M.Boddy@ljmu.ac.uk.

**Abstract**

*Objectives:* To examine the backward comparability of a range of wrist-worn accelerometer estimates of sedentary time (ST) with ActiGraph 100 count∙min-1 waist ST estimates.

*Design:* Cross-sectional, secondary data analysis

*Method:* One hundred and eight 10-11-year-old children (65 girls) wore an ActiGraph GT3X+ accelerometer (AG) on their waist and a GENEActiv accelerometer (GA) on their non-dominant wrist for seven days. GA ST data were classified using a range of thresholds from 23-56 m*g.* ST estimates were compared to AG ST 100 count∙min-1 data. Agreement between the AG and GA thresholds was examined using Cronbach’s alpha, intraclass correlation coefficients (ICC), limits of agreement (LOA), Kappa values, percent agreement, mean absolute percent error (MAPE) and equivalency analysis.

*Results:* Mean AG total ST was 492.4 minutes over the measurement period. Kappa values ranged from 0.31-0.39. Percent agreement ranged from 68-69.9%. Cronbach’s alpha values ranged from 0.88-0.93. ICCs ranged from 0.59-0.86. LOA were wide for all comparisons. Only the 34 m*g* threshold produced estimates that were equivalent at the group level to the AG ST 100 count∙min-1 data though sensitivity and specificity values of ~64% and ~74% respectively were observed.

*Conclusions:* Wrist-based estimates of ST generated using the 34 m*g* threshold are comparable with those derived from the AG waist mounted 100 count∙min-1 threshold at the group level. The 34 m*g* threshold could be applied to allow group-level comparisons of ST with evidence generated using the ActiGraph 100 count∙min-1 method though it is important to consider the observed sensitivity and specificity results when interpreting findings.

*Keywords:* accelerometry, physical activity, sedentary behaviour, children, raw acceleration signals, measurement

**Introduction**

Sedentary behaviour (SB) has received increased attention across recent years as a behaviour that may detrimentally affect children’s health. Whether SB influences health independent of physical activity (PA) is deemed to be a controversial topic, with some studies demonstrating the negative effects of reallocating moderate-to vigorous PA (MVPA) to SB 1, 2, and others reporting limited evidence that SB is associated with health independent of MVPA 3. Nonetheless, researchers are interested in measuring youth movement behaviours including SB to explore health associations, investigate secular trends, and establish intervention effects.

Sedentary behaviour is defined as any waking behaviour characterised by an energy expenditure of ≤1.5 metabolic equivalents (METs) while in a sitting, reclining or lying posture 4. Despite the SB definition referring to posture, many researchers use accelerometers to classify SB as an absence of, or little, registered dynamic acceleration 5. While widely used, it should be noted that this approach does not consider posture.

Historically children’s SB or sedentary time (ST: the time spent for any duration or in any context in sedentary behaviours 4) was assessed using waist-worn ActiGraph accelerometers with a threshold of ≤100 vertical axis count∙min-1 used as the upper boundary for ST. This approach has demonstrated acceptable agreement with measures that classify posture such as the activPAL 6 and those that provide an estimate of energy expenditure, for example indirect calorimetry7. Recently the field has moved towards that of wrist accelerometry due to superior wear compliance 8. Despite observing better compliance, and moderate-to-strong correlations between acceleration data collected using wrist GENEActiv and waist ActiGraph accelerometer placements9, wrist placements generally result in higher estimates of physical activity, therefore wrist and hip data are not directly comparable without correcting for these differences 8. Researchers have also begun to make use of raw acceleration signal analysis to remove the proprietary nature of counts-based data and improve comparability between different devices 10. Although this is advantageous for studies moving forward, a wealth of SB data exists using the ≤100 count∙min-1 threshold applied to data from hip-worn ActiGraphs (for example: 11, 1). Therefore, the ability to compare new raw acceleration derived estimates with those generated using the ≤100 count∙min-1 method would be useful for researchers in the field.

In a previous study authors generated and cross-validated a GENEActiv (GA) wrist threshold of 51 m*g* with the intention of providing comparable estimates of ST to those generate using the traditional ActiGraph waist-worn ≤100 count∙min-1 method. The open source R package GGIR was used to calculate average magnitude of dynamic acceleration, known as the Euclidean Norm Minus One (ENMO), from raw acceleration data12, applying the newly generated (51 m*g*) and published thresholds for SB 13, 14. The comparability of ST estimates between the newly generated 51 m*g* threshold, the other empirical GA raw acceleration thresholds, and ActiGraph waist-worn ≤100 count∙min-1 data was examined. Results demonstrated a lack of equivalence for the 51 m*g* threshold and existing GENEActiv empirical wrist ST thresholds 15, 16. The study provided some preliminary evidence of group-level agreement for the 36 m*g* empirical threshold13 which was originally intended to classify wrist-worn ActiGraph data. A study limitation, however, was that the individual level agreement was undetermined. At the study conclusion, the authors called for the backwards compatibility of ST estimates to be examined further, both at the individual *and* group levels covering a broad range of ENMO thresholds between the lowest (23 m*g*) and highest (56 m*g*) thresholds used to date by researchers in the field. This would enable researchers to establish the most comparable threshold to use when comparing to earlier estimates of ST from ActiGraph data. The lack of evidence related to the comparability of ST estimates currently presents a challenge for researchers when attempting to compare data to those previously recorded using hip and count methods. More investigation is required, therefore, to confirm whether the 36 m*g* proposed in the previous study represents the optimal threshold to use for this purpose.

The backwards comparability of wrist generated ENMO assessed MVPA with traditional accelerometer counts-based data using a range of waist-worn ActiGraph MVPA thresholds has been recently demonstrated. The study proposed wrist ENMO thresholds that gave estimates of MVPA that are comparable with waist measured counts-based data classified using empirical ActiGraph thresholds 17. To date, the backwards comparability for the range of ENMO-derived sedentary behaviour/time estimates has not been comprehensively examined. The aim of this secondary data analysis was therefore to extend previous work by using a wide range of both empirical and arbitrary ST thresholds to examine the backwards comparability of wrist-worn accelerometer estimates of ST with waist-worn ActiGraph 100 count∙min-1 ST. The study also develops previous work by investigating the extent of backwards compatibility at both the individual and group levels.

**Methods**

This is a secondary data analysis, and the methods for the study have been previously published elsewhere 8, 12, 15. Briefly, after gaining institutional ethical approval, parental/carer consent and child assent, 108 10-11-year-old children were involved in this study (65 girls). Body mass was assessed to the nearest 0.1 kg (Seca, Birmingham, UK) and stature was assessed to the nearest 0.1 cm using a portable height meter (Leicester Height Measure, Seca, Birmingham, UK) during school-based data collection sessions conducted between January - May 2014.

Two tri-axial accelerometers (GENEActiv; Activinsights, Cambs, Uk and ActiGraph GT3X+; ActiGraph, Pensacola, FL) were used to assess sedentary time. The GENEActiv (GA) was worn on the non-dominant wrist, and the ActiGraph (AG) worn on the right hip for waking hours over seven consecutive days. Participants were instructed to remove the monitors when engaging in water-based activities and also when sleeping. Both monitors were initialised to record at a sampling frequency of 100 Hz using the same computer.

Data generated by the AG devices were processed using ActiLife v 6.11.4 software (ActiGraph, Pensacola, FL). Consistent with previous research 15, 18, for the AG devices, non-wear was defined as 20 minutes of consecutive zero counts (1-minute spike tolerance) and was subtracted from daily wear time. A valid day was defined as ≥540 min for weekdays 19 and ≥480 min for weekend days 20. Consistent with a previous study 15, the valid weekend and weekdays with the longest wear for each participant were retained for analysis. Where participants did not have a valid weekend day, only their longest valid weekday was retained for analysis. Data for the included days were converted into 1 second csv files, with non-wear times manually removed at a later step in data reduction described later. Sedentary time was defined as ≤100 count∙min-1 21 and coded accordingly. Data generated by the GA monitors were saved as binary files after being downloaded using GENEActiv v 2.2 software (Activinsights, Cambs, UK). GA data were processed in R using the GGIR package version 1.1-4 to calculate the ENMO-derived average magnitude of dynamic acceleration. ENMO is vulnerable to calibration errors, therefore to correct for sensor calibration errors, autocalibration was completed 22. Non-wear for the GA data was scored using 60 minutes moving windows with 15 minutes increments and imputation was completed 23. ENMO values were expressed in average m*g* per 1 second epoch23, and GA data for the corresponding AG week and/or weekend days were retained for further analysis.

Time stamps for the GA and AG were synchronised, and data were merged resulting in one csv file for each participant. Periods of non-wear were manually removed from both the AG and GA data according to the wear details generated by the ActiLife AG analysis. Therefore all epochs remaining in the dataset contained ‘wear’ data for both devices. After non-wear periods were removed, data were then reduced to 1-minute epochs and AG data were scored as sedentary or active using vertical axis 100 count∙min-1 as the reference value for ST 21. In the previous study ST thresholds of 23 m*g* (obtained by solving the Hildebrand et al., (2014) regression equation), 36 m*g* 13, 51 m*g* (newly generated and cross-validated threshold), and 56 m*g* 13 were used. In Step 1 of the analysis for the current study we extend these results to examine comparability of GA ST data classified using a wider range of thresholds. This included a recently published ST threshold of 52 m*g* that was generated by a child-specific calibration circuit24 and arbitrary thresholds of 30 m*g*, 40 m*g* and 45 m*g* which were chosen to cover the range of thresholds. The final thresholds included in Step 1 were therefore: 23 m*g*, 30 m*g*, 36 m*g*, 40 m*g*, 45 m*g*, 52 m*g* and 56 m*g*. This approach resulted in a range of ST thresholds with which to compare to the AG vertical axis 100 count∙min-1 reference.

Following calculation of descriptive statistics to describe the participant group, the GA ST estimates were compared to the AG 100 count∙min-1 estimates at the group level by calculating Cronbach’s alpha, intraclass correlation coefficients (ICC), Kappa values and percent agreement. Individual level estimates were compared by calculating limits of agreement (LOA), correlations between bias and mean sedentary time (AG and GA) and mean absolute percent error (MAPE, %). Null hypothesis testing is not appropriate when considering the comparability between estimates 25, therefore equivalency analysis was also performed to establish the equivalence of group level estimates of ST on average. A 95%equivalence test was completed to establish whether the 90% confidence intervals for the range of GA ST thresholds completely fell within the zone of equivalence, defined as ±10% of the mean AG 100 count∙min-1 classified ST.

In Step 2 of the analysis, results from step 1 were used to identify the likely range of most comparable thresholds, and further thresholds within this range were then added to the analysis to attempt to find the optimal threshold. Analyses were also completed separately by sex to further examine the comparability of estimates. To provide a more stringent comparison, for the second step in analysis, the zone of equivalence for the group-level equivalence test was reduced to ±5% of the mean AG 100 count∙min-1 classified ST. Analysis was completed using IBM SPSS Statistics v.24 (IBM, Armonk, NY) and Microsoft Excel 2016 (Microsoft, Redmond, WA).

**Results**

Participant characteristics, the number of days included in analyses, weekday and weekend day wear times for boys and girls have been published previously for this population 15, and are displayed in Table 1.

[TABLE 1 ABOUT HERE]

Analysis Step 1. Table 2 summarises the comparisons between the AG ≤100 count∙min-1 and various GA ST estimates. Kappa values ranged from 0.31-0.39, representing ‘fair’ agreement 26. Percent agreement ranged from 68-69.9%. Cronbach’s alpha values were 0.88 for the 23 m*g* threshold, suggesting a good level of consistency, where all other Cronbach’s alpha values were >0.9, suggesting excellent levels of consistency. ICCs ranged from 0.59 for the 23 m*g* threshold (moderate reliability) to 0.86 for the 36 m*g* threshold (good reliability). Supplementary content A displays the Bland-Altman plots for the comparisons. LOA were wide for AG - GA comparisons, with the narrowest limits observed for the AG v 36 m*g* comparison (lower LOA = -230.47 upper LOA = 194.81 minutes), with systematic bias apparent. All thresholds from 36 m*g* and above showed negative bias illustrated by mean bias and the correlation results (i.e. higher GA ST estimates than AG). The highest negative bias observed for the 56 m*g* threshold. MAPE (%) ranged from 15.8% for the 36 m*g* threshold to 40.7% for the 56 m*g*. The results of the equivalency analysis found that only the ST estimates generated by the 36 m*g* threshold could be considered statistically equivalent to the AG ≤100 count∙min-1 on average at the group level with 90% CI’s falling completely within the ±10% zone of equivalence. Thresholds ≤30 m*g* appeared to underestimate and ≥40 m*g* appeared to overestimate ST in comparison to the ST reported using AG ≤100 count∙min-1. Therefore, for analysis Step 2 thresholds of 34 m*g* and 35 m*g* were included to examine the optimum threshold and analyses were repeated for the whole cohort and separately by sex. The zone of equivalence was reduced to ±5% for the group-level equivalency analysis. Table 3 summarises the comparisons between the AG ≤100 count∙min-1 and 34 m*g* and 35 m*g* GA ST estimates and includes sensitivity and specificity information. Mean bias was low for the 34 m*g* threshold though wide limits of agreement were observed for both thresholds and MAPE% was similar to that observed for the 36 m*g* threshold at 16.2% (34 m*g*) and 15.8% (35 m*g*). Sensitivity values (true positive) were similar between the thresholds, at 63.6% and 64.8% for 34 mg and 35 mg respectively. Specificity values (true negative) were also similar, at 74.2% and 73.4% for the 34 m*g* and 35 m*g* thresholds respectively. Boys’ data displayed wider limits of agreement, higher MAPE% and slightly higher sensitivity values for both thresholds in comparison to girls, though % agreement, Cronbach’s alpha, ICC, Kappa and and specificity values were similar. The results of the equivalence analysis for all threshold comparisons are displayed in Figure 1. Only the ST estimates generated by the 34 m*g* threshold were statistically equivalent to the AG ≤100 count∙min-1 on average at the group level with 90% CI’s falling completely within the ±5% zone of equivalence.

[TABLE 2 ABOUT HERE]

[TABLE 3 ABOUT HERE]

[FIGURE 1 ABOUT HERE]

**Discussion**

A wealth of existing accelerometer data has used the threshold of ≤100 count∙min-1 applied to waist-worn ActiGraphs to determine time spent sedentary (for example large studies using the International Children’s Accelerometry Database 11, 1). As the discipline moves increasingly towards raw acceleration data processing and wrist-worn monitors, the ability for researchers to compare data between studies that have used counts-based processing methods and waist-worn monitors is important. The aim of this study was therefore to examine the backwards comparability of wrist-worn accelerometer estimates of sedentary time (ST) with ActiGraph 100 count∙min-1 waist ST estimates using a range of empirically determined and arbitrary raw acceleration thresholds for wrist-worn monitors.

This study has demonstrated moderate to excellent ICC values, and moderate to good Cronbach’s alpha values at the group level for all the GA thresholds. Despite these results, Kappa values were ‘fair’ and large MAPE values (individual level) were observed. In addition, wide limits of agreement (individual level) were observed between all GA thresholds and the AG standard. Systematic bias was evident, indicating that as estimates of ST increased so did the bias. Equivalency analysis found no thresholds produced estimates of ST that could be considered statistically equivalent on average at the group level in comparison to the AG standard with the exception of the 34 m*g* threshold. The wide limits of agreement, MAPE and bias results, in the presence of high consistency, as evidence by high ICC and Cronbach’s alpha, highlights the importance of considering a range of analyses at the individual and group levels when examining the comparability of ST estimates.

Our previous work called for studies to investigate the backwards compatibility of ST estimates 15. In the present study this was addressed by examining a broad range of thresholds, both empirically determined and arbitrary, and out of the selected thresholds the 34 m*g* threshold provided the ST estimates most comparable to the ActiGraph 100 count∙min-1 waist ST at the group level. Despite this, the limits of agreement showed that ST estimates generate using the 34 mg threshold ranged from ±~-3 hrs in comparison to the AG estimates, therefore suggesting the 34 m*g* should not be used for individual level comparison. In our previous study we established that the 36 m*g* threshold, provided equivalent estimates of wrist ST for the GENEActiv as the ≤100 count∙min-1 standard at the group level when using a ±10% zone of equivalence. The present study that used a more stringent ±5% zone of equivalence suggests that 34 m*g* may provide a more accurate comparison at the group level, and furthermore suggests that lower and higher thresholds across the range are not appropriate for this purpose. ActiGraph accelerometers are known to produce lower ENMO values than GA devices 8, though recent evidence suggests the GA and AG devices provide equivalent estimates between the 30-50 m*g* range 16. Irrespective of potential differences between devices, at the group level the 34 m*g* threshold provided the most comparable estimates of ST to the AG hip ≤100 count∙min-1 standard, so could be used for comparative purposes across studies moving forward.

Despite exhibiting group level equivalency, the sensitivity and specificity values suggest that for every 100 minutes of ST classified by the ActiGraph, the GA 34 m*g* threshold would classify ~64 minutes of ST. Therefore any comparisons between studies using the wrist worn 34 m*g* threshold and studies using the waist worn AG ≤100 count∙min-1 method should bear the sensitivity and specificity results in mind when interpreting findings. Furthermore, the 5% zone of equivalence provides a range of ~50 minutes of sedentary time which the 34 mg estimates fell within. Whether a potential difference of ±~50 minutes is clinically meaningful or whether that would provide estimates that are sensitive to change is open to debate. Recent evidence suggests that the reallocation of 15 minutes of sedentary time to moderate to vigorous physical activity (MVPA) predicted changes in obesity and fitness outcomes in children2. However, such evidence relies on sedentary time to be reallocated to MVPA, and the impact of reallocation of time to light intensity physical activity or stationary behaviours independent of MVPA remains unclear. Where group-level comparisons with data collected using the AG hip ≤100 count∙min-1 standard are useful the 34 m*g* threshold can be applied, though where precise estimates of sedentary time or behaviour are required to demonstrate intervention effectiveness or individual level changes alternative methods may be required. Whether the estimates of ST from the GA and AG reflect actual ST remains open to debate. Indeed it is questionable whether the absence or low levels of acceleration should be used in isolation to classify ST, especially considering the postural component that is integral to the definition of sedentary behaviour. Examining the accuracy of measuring ST was not the aim of the present study *per se*, and as such represents a different research question to be addressed in the future. There are, however, ways of processing accelerometer data to classify posture that do not require the use of additional devices or monitoring periods. One example is the sedentary sphere 27, which classifies assumed postural changes based on acceleration signals, arm orientation and wrist orientation. Although this approach has shown promise in adult populations, it has not yet been validated in children and so its utility in this population has not been established. Nonetheless, estimates based on new approaches, irrespective of the method, still raises questions regarding the comparability with the large volume of existing literature therefore a pragmatic solution is warranted.

There are some limitations to the present study. We used a 1-minute epoch to determine time spent sedentary to allow a comparison to the AG hip ≤100 count∙min-1 standard. The majority of children’s ST data using AG hip ≤100 count∙min-1 utilises 1-minute epochs, therefore this approach was necessary to address the study aims. It is well established that children’s physical activity behaviours are sporadic in nature 28, 29 and though high frequency monitoring is required to detect movement at higher intensities, the 1 minute epoch is unlikely to influence recorded ST which is generally accrued in bouts lasting >2 minutes 30. In addition, the group of participants involved in this study were all from the same geographical location in North-West England and a narrow age-range, therefore their ST behaviours may not be representative of different populations and groups. We included a maximum of 2 days of data (one weekend and one weekday) for each participant, therefore the sedentary levels of participants are not reflective of their habitual patterns. However, the volume of data included allows for comparison between devices and signal classification, where 7 day’s data would have been prohibitive in terms of file size. Furthermore, a waking hours accelerometer protocol was used. Therefore recent studies using 24-hour protocols may require further investigation to examine the backward compatibility of data, including sleep classification in addition to ST estimates.

To the best of our knowledge, this is the first study to examine the backward comparability of wrist assessed sedentary time with ActiGraph 100 count∙min-1 waist ST estimates. The results of the study suggest that the 34 m*g* threshold produced the most comparable estimates of ST and could be used to classify data for group-level comparison with previously published studies that used the 100 count∙min-1 threshold.

**Conclusions**

Despite observing high ICC and Cronbach’s alpha values, the results suggest that the all but one of the wrist mounted, raw acceleration derived ST estimates should not be directly compared with those derived from the 100 count∙min-1 waist mounted AG threshold. The 34 m*g* threshold may provide comparable ST estimates at the group level, and future studies could use the 34 m*g* threshold when comparing ENMO derived ST estimates group level estimates previously published using the 100 count∙min-1 approach though it is important to consider the observed sensitivity and specificity results when interpreting findings.

**Practical Implications**

* Many previous studies estimated children’s sedentary time using waist-mounted ActiGraph accelerometers and the 100 count∙min-1 threshold.
* The backward comparability of wrist-worn raw acceleration derived sedentary time estimates with the wealth of data collected using waist-mounted ActiGraphs is unknown.
* This study found that the 34 m*g* threshold could be applied to wrist accelerometer data to provide estimates of sedentary time that are equivalent to the ActiGraph waist-worn 100 count∙min-1 on average at the group level, though the sensitivity and specificity values observed in this study should be considered when interpreting findings.

**Acknowledgements**

The authors would like to thank the children and schools that were involved this research. AR is with the National Institute for Health Research (NIHR) Biomedical Research Centre based at University Hospitals of Leicester and Loughborough University, the National Institute for Health Research Collaboration for Leadership in Applied Health Research and Care – East Midlands (NIHR CLAHRC – EM) and the Leicester Clinical Trials Unit. The views expressed are those of the authors and not necessarily those of the NHS, the NIHR or the Department of Health.

**References**

1. Hansen BH, Anderssen SA, Andersen LB, et al. Cross-Sectional Associations of Reallocating Time Between Sedentary and Active Behaviours on Cardiometabolic Risk Factors in Young People: An International Children's Accelerometry Database (ICAD) Analysis. *Sports Med.* 2018; 48(10):2401-2412.

2. Fairclough SJ, Dumuid D, Mackintosh KA, et al. Adiposity, fitness, health-related quality of life and the reallocation of time between children's school day behaviours: A compositional data analysis. *Preventive Medicine Reports.* 2018; 11:254-261.

3. Cliff DP, Hesketh KD, Vella SA, et al. Objectively measured sedentary behaviour and health and development in children and adolescents: systematic review and meta-analysis. *Obes Rev.* 2016; 17(4):330-344.

4. Tremblay MS, Aubert S, Barnes JD, et al. Sedentary behavior network (SBRN)- Terminology Consensus Project process and outcome. *International Journal of Behavioural Nutrition and Physical Activity.* 2017; 14(75).

5. Lubans DR, Hesketh K, Cliff DP, et al. A systematic review of the validity and reliability of sedentary behaviour measures used with children and adolescents. *Obes Rev.* 2011; 12(10):781-799.

6. Ridgers ND, Salmon J, Ridley K, E. OC, Arundell L, Timperio A. Agreement beween activPAL and ActiGraph for assessing children's sedentary time. *Int J Behav Nutr Phys Act.* 2012; 9(1):15.

7. Evenson KR, Catellier DJ, Gill K, Ondrak KS, McMurray RG. Calibration of two objective measures of physical activity for children. *J Sports Sci.* 2008; 26(14):1557-1565.

8. Fairclough SJ, Noonan R, Rowlands AV, Van Hees V, Knowles Z, Boddy LM. Wear Compliance and Activity in Children Wearing Wrist- and Hip-Mounted Accelerometers. *Med Sci Sports Exerc.* 2016; 48(2):245-253.

9. Rowlands AV, Fraysse F, Catt M, et al. Comparability of measured acceleration from accelerometry-based activity monitors. *Med Sci Sports Exerc.* 2015; 47(1):201-210.

10. Rowlands AV, Olds TS, Hillsdon M, et al. Assessing sedentary behavior with the GENEActiv: introducing the sedentary sphere. *Med Sci Sports Exerc.* 2014; 46(6):1235-1247.

11. Cooper AR, Goodman A, Page AS, et al. Objectively measured physical activity and sedentary time in youth: the International children's accelerometry database (ICAD). *Int J Behav Nutr Phys Act.* 2015; 12:113.

12. Noonan RJ, Boddy LM, Kim Y, Knowles ZR, Fairclough SJ. Comparison of children's free-living physical activity derived from wrist and hip raw accelerations during the segmented week. *J Sports Sci.* 2017; 35(21):2067-2072.

13. Hildebrand M, Hansen BH, van Hees VT, Ekelund U. Evaluation of raw acceleration sedentary thresholds in children and adults. *Scand J Med Sci Sports.* 2016.

14. Hildebrand M, VT VANH, Hansen BH, Ekelund U. Age group comparability of raw accelerometer output from wrist- and hip-worn monitors. *Med Sci Sports Exerc.* 2014; 46(9):1816-1824.

15. Boddy LM, Noonan RJ, Kim Y, et al. Comparability of children's sedentary time estimates derived from wrist worn GENEActiv and hip worn ActiGraph accelerometer thresholds. *Journal of science and medicine in sport / Sports Medicine Australia.* 2018.

16. Rowlands AV, Mirkes EM, Yates T, et al. Accelerometer-assessed Physical Activity in Epidemiology: Are Monitors Equivalent? *Med Sci Sports Exerc.* 2018; 50(2):257-265.

17. Rowlands AV, Cliff DP, Fairclough SJ, et al. Moving Forward with Backward Compatibility: Translating Wrist Accelerometer Data. *Med Sci Sports Exerc.* 2016; 48(11):2142-2149.

18. Catellier DJ, Hannan PJ, Murray DM, et al. Inputation of missing data when measuring activity by accelerometry. *Med Sci Sports Exerc.* 2005; 37(Suppl 11):S555-S562.

19. Rich C, Geraci M, Griffiths L, Sera F, Dezateux C, Cortina-Borja M. Quality control methods in accelerometer data processing: defining minimum wear time. *PloS one.* 2013; 8(6):e67206.

20. Rowlands AV, Pilgrim EL, Eston RG. Patterns of habitual activity across weekdays and weekend days in 9-11-year-old children. *Prev Med.* 2008; 46(4):317-324.

21. Trost SG, Loprinzi PD, Moore R, Pfeiffer KA. Comparison of accelerometer cut points for predicting activity intensity in youth. *Med Sci Sports Exerc.* 2011; 43(7):1360-1368.

22. van Hees VT, Fang Z, Langford J, et al. Autocalibration of accelerometer data for free-living physical activity assessment using local gravity and temperature: an evaluation on four continents. *Journal of applied physiology.* 2014; 117(7):738-744.

23. van Hees VT, Gorzelniak L, Dean Leon EC, et al. Separating movement and gravity components in an acceleration signal and implications for the assessment of human daily physical activity. *PLoS One.* 2013; 8(4):e61691.

24. Hurter L, Fairclough SJ, Knowles ZR, Porcellato LA, Cooper-Ryan AM, Boddy LM. Establishing Raw Acceleration Thresholds to Classify Sedentary and Stationary Behaviour in Children. *Children (Basel).* 2018; 5(12).

25. Dixon PM, Saint-Maurice PF, Kim Y, Hibbing P, Bai Y, Welk GJ. A primer on the use of equivalence testing for evaluating measurement agreement. *Medicine and Science in Sports and Exercise.* 2017; Ahead of Print.

26. Altman D. *Practical Statistics for Medical Research*, London, Chapman and Hall; 1991.

27. Rowlands AV, Yates T, Olds TS, Davies M, Khunti K, Edwardson CL. Sedentary Sphere: Wrist-Worn Accelerometer-Brand Independent Posture Classification. *Med Sci Sports Exerc.* 2016; 48(4):748-754.

28. Bailey RC, Olson J, Pepper SL, Porszasz J, Barstow TJ, Cooper DM. The level and tempo of children's physical activities: an observational study. *Med Sci Sports Exerc.* 1995; 27(7):1033-1041.

29. Baquet G, Stratton G, Van Praagh E, Berthoin S. Improving physical activity assessment in prepubertal children with high-frequency accelerometry monitoring: a methodological issue. *Prev Med.* 2007; 44(2):143-147.

30. Verloigne M, Ridgers ND, Chinapaw M, et al. Patterns of objectively measured sedentary time in 10- to 12-year-old Belgian children: an observational study within the ENERGY-project. *BMC Pediatr.* 2017; 17(1):147.

Table and Figure Legends

Table 1. Mean (SD) anthropometric, wear time and number of days included within analysis for boys and girls

|  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- |
|  | Boys N = 43 | |  | Girls N = 65 | |
|  | Mean or Frequency | SD |  | Mean or Frequency | SD |
| Age (y) | 10 | 0.4 |  | 10 | 0.3 |
| Height (cm) | 139.5 | 7.9 |  | 138 | 7.4 |
| Body mass (kg) | 35.6 | 8.2 |  | 34.2 | 8.6 |
| BMI (kg∙m.2) | 18.2 | 3.00 |  | 17.8 | 3.2 |
| ActiGraph weekday wear (min∙day-1) | 739.9 | 115.6 |  | 738.8 | 100.4 |
| ActiGraph weekend day wear (min∙day-1) | 631.8 | 110.8 |  | 661.5 | 108.3 |
| ActiGraph valid weekdays included | 41 | N/A |  | 64 | N/A |
| ActiGraph valid weekend days included | 30 | N/A |  | 46 | N/A |
| Total valid included days | 71 | N/A |  | 110 | N/A |

Table 2. Comparisons between the ActiGraph ≤100 count∙min-1 standard and GA thresholds

|  |  |  |  |  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- |
| Criterion | GA threshold | Sedentary time (mins/included days) | Mean Bias (mins) | Bias vs Mean ST correlation | Cronbach’s Alpha | ICC (Single measures) | Limits of agreement Lower | Limits of Agreement upper | Mean absolute percent error % (SD) | % agreement | Kappa |
| ActiGraph ≤100 count∙min-1 |  | 492.4 |  |  |  |  |  |  |  |  |  |
|  | 23m*g* | 342.6 | 149.8 | -.11 | 0.88 | 0.59 | -88.38 | 387.94 | 35.5 (18.2) | 68 | 0.31 |
|  | 30m*g* | 440.4 | 52 | -.33\*\* | 0.91 | 0.82 | -166.41 | 270.45 | 20.2 (16.5) | 69.4 | 0.36 |
|  | 36m*g* | 510.2 | -17.8 | -.49\*\* | 0.93 | 0.86 | -230.47 | 194.81 | 15.8 (15.7) | 69.8 | 0.38 |
|  | 40m*g* | 554.8 | -62.4 | -.58\*\* | 0.93 | 0.83 | -276.22 | 151.36 | 18.2 (16.3) | 69.9 | 0.39 |
|  | 45m*g* | 603.6 | -111.2 | -.67\*\* | 0.93 | 0.77 | -327.29 | 104.95 | 25.1 (16.4) | 69.5 | 0.39 |
|  | 52m*g* | 660 | -167.6 | -.75\*\* | 0.93 | 0.69 | -391.09 | 55.91 | 34.9 (16.7) | 68.8 | 0.39 |
|  | 56m*g* | 692.3 | -199.9 | -.79\*\* | 0.93 | 0.63 | -430.48 | 30.7 | 40.7 (18) | 68.2 | 0.38 |

Table 3. Comparisons between the ActiGraph ≤100 count∙min-1 standard and 34 m*g* and 35 m*g* GA thresholds.

|  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- |
| Criterion | GA threshold | Sedentary time (mins) | Mean Bias (mins) | Bias vs Mean ST correlation | Cronbach’s Alpha | ICC (Single measures) | Limits of agreement Lower | Limits of Agreement upper | Mean absolute percent error % (SD) | % agreement | Kappa | Sensitivity % | Specificity % |
| ActiGraph ≤100 count∙min-1 |  | 492.4 |  |  |  |  |  |  |  |  |  |  |  |
|  | 34 m*g* | 489.6 | 2.8 | -.38\*\* | 0.92 | 0.86 | -219.51 | 216.18 | 16.2 (15.8) | 69.8 | 0.38 | 63.6 | 74.2 |
|  | 35 m*g* | 501 | -8.6 | -.47\*\* | 0.93 | 0.86 | -220.88 | 203.71 | 15.8 (15.7) | 69.8 | 0.38 | 64.8 | 73.4 |
| Boys |  |  |  |  |  |  |  |  |  |  |  |  |  |
| ActiGraph ≤100 count∙min-1 |  | 499.7 |  |  |  |  |  |  |  |  |  |  |  |
|  | 34 m*g* | 494.7 | 5 | -.47\*\* | 0.93 | 0.87 | -236.3 | 246.3 | 19 (16.8) | 70.4 | 0.40 | 65.6 | 74.2 |
|  | 35 m*g* | 505.5 | -5.8 | -.52\*\* | 0.93 | 0.87 | -247.3 | 235.6 | 19 (16.2) | 70.4 | 0.40 | 66.7 | 73.3 |
| Girls |  |  |  |  |  |  |  |  |  |  |  |  |  |
| ActiGraph ≤100 count∙min-1 |  | 487.6 |  |  |  |  |  |  |  |  |  |  |  |
|  | 34 m*g* | 486.1 | 1.4 | -.31\*\* | 0.92 | 0.85 | -193.2 | 196 | 14.4 (14.9) | 69.4 | 0.37 | 62.3 | 74.2 |
|  | 35 m*g* | 498 | -10.4 | -.42\*\* | 0.92 | 0.86 | -202.9 | 182.1 | 13.7 (15) | 69.4 | 0.37 | 63.6 | 73.5 |

Figure 1. ActiGraph ≤100 count∙min-1 ±5% zone of equivalence (467.7 minutes- 517 minutes, dotted lines) and 90% confidence intervals for the GENEActiv sedentary time estimates classified using nine thresholds