1 The Geography of Talent in China During 2000-2015: An Eigenvector Spatial

2 Filtering Negative Binomial Panel Data Approach

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4 Abstract: The increase in China's skilled labor force has drawn much attention from 5 policymakers, national and international firms and media. Understanding how educated 6 talent locates and re-locates across the country can guide future policy discussions of equality, firm localization and service allocation. Prior studies have tended to adopt a 7 8 static cross-national approach providing valuable insights into the relative importance 9 of economic and amenity differentials driving the distribution of talent in China. Yet, how the influence of these factors has evolved over time remains unexplored. Recently 10 released official statistical data now enables space-time analysis of the geographic 11 12 distribution of talent and its determinants in China. Using four-year city-level panel 13 data from national population censuses and 1% population sample surveys conducted 14 every five years between 2000 and 2015, we examine the spatial patterns of talent 15 across Chinese cities and their underpinning drivers evolve over time. Our results 16 revealed that the spatial distribution of talent in China is persistently unequal and 17 spatially concentrated between 2000 and 2015, although the extent of this unevenness and concentration has decreased slightly over time. Results also showed that economic 18 opportunities have remained main attraction and retention forces of talent in main cities, 19 20 but that key local amenities have also played a role in shaping the residential decisions of talent, particularly urban public services and greening rate. These results highlight 21 22 that China's talent's settlement patterns over the first fifteen years of the 21st century 23 have been mainly driven by urban economic opportunities versus amenity-related 24 factors.

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Keywords: talent; spatial pattern; determinants; eigenvector spatial filtering; panel data
 analysis; China

28

29 **1. Introduction**

30 Human capital accumulation has been recognized as a vital driving force in the 31 development of knowledge-based economies (Lucas, 1988; Romer, 1990). As a result, policymakers have been universally concerned with ways of attracting and retaining 32 33 human talent (Florida, 2002). In China, the national number and share of the highly 34 educated population as a share of the labor force have both rapidly increased over the 35 last fifteen years. The 2015 national 1% population sample survey reported that 170.93 million Chinese citizens held a college degree or above, an increase of 72.4% from the 36 same sample collected in 2000. The rise of skilled individuals in China's labor force 37 38 has drawn much attention from policymakers, employers, and the media. China's 39 government has placed increasing emphasis on enlarging the national pool of highly skilled laborers by developing The National Medium and Long-Term Talent 40 Development Plan (2010-2020) in 2010 (State Council of China, 2010). This plan states 41 42 that China needs to transition from labor-intensive to talent-intensive economic

43 activities to increase its competitiveness in the global economy. Similarly, the report of 44 the 19th National Congress of China in 2017 pointed out that China should firmly 45 implement "the strategy of reinvigorating China through human resource development". It also states that China should cultivate a large number of strategic 46 talent, leading talent in the fields of science and technology, young talent, and highly 47 48 innovative teams with international standards. Over the last 10 years, local governments 49 at all levels have been engaged in the competition for skilled workers on the global 50 stage, implementing policies to attract and retain talent, such as the "Peacock Plan" 51 (kong que ji hua) in Shenzhen (Shenzhen Municipal Government, 2019). These policies 52 include housing allowances, fast-track hukou transfers, and lower thresholds for bank 53 loan applications.

54 There has been a surge in the number of studies on the migration and redistribution 55 of various kinds of talent in China over the past decade (Florida et al., 2012; Gu et al., 2019a; Liu & Shen, 2014a, 2014b; Liu et al., 2017; Liu & Xu, 2017; Nie & Liu, 2018; 56 57 Qian, 2010; Zeng et al., 2019). Due to data availability, prior studies have tended to 58 adopt a static cross-national approach providing novel insights into the relative 59 importance of economic and amenity differentials driving the migration patterns of talent. Yet, changes in the influence of these factors on shaping shifts in the spatial 60 dynamics of migration over time remain underexplored. Understanding these changes 61 is key to guide evidence-based economic development policies targeted at attracting 62 63 and retaining talent. Additionally, the geographical scale of existing studies has been 64 relatively coarse (i.e. provincial level), restricting our understanding of the location of 65 talent across the urban-rural continuum within and between provinces. Moreover, prior empirical studies tend to focused on understanding the migration and redistribution 66 patterns of highly educated talent during the period pre-2010. Yet less is known about 67 68 the contemporary patterns of Chinese highly educated talent post-2010. The availability 69 of the 2010 census and 2015 1% sample survey data enables investigating more recent spatial patterns of talent in China and identifying their key drivers. 70

71 More generally, while previous analysis on migration in Europe has identified a 72 pattern of highly significant spatial autocorrelation in the distribution of talent 73 (Miguélez et al., 2010; Rodríguez-Pose & Tselios, 2011), little has been done to model 74 these spatial effects in a regression framework explicitly. It is shown that talent stock 75 in a particular region may be positively related to the pool of talent in its surrounding 76 areas, which leads to spatial autocorrelation. Researchers have discussed the sources 77 for the spatial autocorrelation: they observed that talent gathered in nearby regions for 78 lower communication cost, knowledge spillover, and sharing public resources (Fujita 79 & Thisse, 1996; Miguélez et al., 2010). Interregional frequent mobility of talent 80 between neighboring areas may also lead to positive spatial autocorrelation in talent 81 stocks (Liu & Shen, 2017), which may also result from close cross-regional social talent networks (Gu et al., 2019a). A handful of studies have used autocorrelation analysis to 82 83 investigate the distribution of talent within a country or a region (Nie & Liu, 2017; Sone 84 et al., 2016). Yet little research has taken into account the presence of spatial

dependence when assessing the geographic distribution of talent within countries. 85 86 Without proper treatment, latent spatial autocorrelation may lead to estimation bias in regression models. Eigenvector spatial filtering (ESF) is a technique for capturing latent 87 88 spatial autocorrelation in model residuals (Liu & Shen, 2017; Gu et al., 2019b). Each 89 eigenvector extracted from a spatial weight matrix can be considered as a control 90 variable for spatial autocorrelation. Hence, incorporating eigenvectors into regression 91 models can help to reduce the effect of spatial autocorrelation, thus reduce model bias 92 and enhancing model fitting (Griffith, 2003; Chun & Griffith, 2011).

93 To address these research gaps, this study aims to examine the spatial-temporal patterns of talent in China from 2000 to 2015 drawing on a unique four-year city-level 94 95 panel data set built from population censuses and 1% population sample surveys. We 96 apply a range of measures of inequality to assess the imbalance degree of the spatial 97 distribution of highly educated talent in China. To tackle any estimation bias arising 98 due to spatial autocorrelation (Griffith, 2003), we build an eigenvector spatial filtering 99 negative binomial panel model (ESF-NBPM) to examine the determinants of the 100 geographical distribution of talent. The contribution of this article is threefold. First, we 101 analyzed the spatial patterns of talent in China at a finer geographical resolution (i.e., the city-level), compared to previous analysis based on provincial-level data. Our more 102 detailed geographic analysis will contribute to the literature on the migration patterns 103 of Chinese talent by developing an understanding of the intra-provincial differences in 104 the locational choices of talent and providing more accurate identification of the key 105 106 underpinning factors of these choices. Second, our analysis also contributes to our 107 understanding of changes in the relative importance of local factors shaping the distribution of talent over time. Third, we applied the ESF specification to tackle spatial 108 109 autocorrelation moving away from traditional gravity models which omit the spatial 110 interconnectivity in the distribution of talent stocks.

The structure of this paper is as follows: Section 2 reviews the existing literature on spatial patterns and determinants of talent. Data sources and research methods are described in Section 3. Sections 4 determines the spatial concentration or evenness in the spatial distribution of talent in China and analyses its evolution over time based on a range of inequality measures. Section 5 constructs an ESF-NBPM to identify the key driving factors of the spatial distribution of talent. The final section summaries the key findings and discusses policy recommendations.

118

119 **2. Literature review**

120 **2.1 Determinants of the geographic distribution of talent**

Economic theories have made considerable efforts to identify the factors shaping the spatial distribution and redistribution of talent (Lewis, 1954; Sjaastad, 1962; Todaro, 1969). New classical theory assumes migration and redistribution as the resulting process of individuals thinking rationally and independently, aiming to maximize their utility (Greenwood, 1975). Empirical results evidence that differentials in regional economic opportunities are dominant factors influencing the spatial distribution of talent. The skilled labor force is spatially focused and thus reinforcing
agglomeration economics (Greenwood, 1975; Harris & Todaro, 1970; Ranis & Fei,
1961). Recent research has shown that highly skilled and educated migrants tend to
move to regions with high-income levels and abundant job opportunities (Arntz, 2010;
Liu & Shen, 2014b; Rowe 2013a; Scott, 2010). Employment stability and job
opportunity also play an important role in attracting skilled labor (Fielding, 1989;
Findlay et al., 2008; Xu & Ouyang, 2018).

134 Alternatively, additional theories and models have been developed which 135 emphasize the role that amenity differentials play in the redistribution of talent. They assume that spatial differences in economic opportunities reflect largely compensating 136 differentials related to corresponding spatial differences in amenities (Graves 1976, 137 138 2018). Based on the principles of Graves' equilibrium model, Glaeser et al. (2001), 139 Clark et al. (2002) and Florida (2002) expanded economic-driven migration models to include urban amenities. Empirical research has shown that strong relationship between 140 regional amenities and redistribution of talent exists, particularly in terms of natural 141 142 environmental amenities (Knapp & Gravest, 1989; Partridge, 2010; Rappaport, 2009), 143 urban public services (Woodward et al., 2006; Rowe et al. 2017), and urban consumer 144 facilities (Glaeser & Gottlieb, 2006).

The location of universities has emerged as a key urban amenity factor displaying a significant relationship with the distribution of highly educated individuals (Florida et al. 2008; Rowe 2013). Universities are key producers of talent. In less-attractive and underdeveloped regions, universities may also play the role of "talent exporters" since local highly educated graduates tend to move to other regions for employment (Florida et al., 2006). When talent is less mobile and has a higher settlement intention of the region, universities have a more notable effect on the pool of local talent (Qian, 2010).

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153 **2.2 Geography of talent in China**

There has been a recent surge in the number of studies examining the spatial 154 155 patterns of talent in China and their underpinning factors (Gu et al., 1999; Liu & Shen, 2014a, 2014b; Liu & Xu, 2017; Nie & Liu, 2018; Yu et al., 2019). Collectively, these 156 157 studies revealed an unbalanced spatial distribution at city or provincial scales (Liu & 158 Shen, 2014a, 2014b; Nie & Liu, 2018). They have also demonstrated the critical role 159 of economic factors influencing the mobility decisions of a range of different types of 160 talent labor, including highly skilled individuals (Miu et al., 2013), scientists (Zhang et al., 2011), and highly educated people (Liu & Shen, 2014a). They have shown that 161 162 economic factors including GDP, average wage, unemployment rate, and industrial 163 structure comprise the fundamental forces shaping the spatial distribution of China's 164 population and talent (Liu & Shen, 2014b; Yu et al., 2019). They have also indicated that the uneven distribution of talent has been the result of a pronounced gap between 165 developed and undeveloped regions in terms of economic development and living 166 167 conditions (Liu & Shen, 2014a). Gu et al. (1999) revealed that economic factors, such 168 as job opportunities and living conditions, are the main drivers of population

distribution. Furthermore, the siphoning effects of talent spurring economic
development in major cities may negatively impact small and medium-sized
settlements, reinforcing spatial economic inequalities.

Consistent with Graves (1976, 2018) and Glaeser (2001), scholars have explored 172 the role of amenities in influencing the distribution of talent in China (Deng et al., 2016; 173 174 Liu & Shen, 2014b). Evidence indicates that public services, such as education and 175 medical care represent a necessary condition to attract talent (Liu & Shen, 2014b). The 176 local share of talent is also correlated with the local quality of the living and 177 consumption facilities, including leisure and entertainment infrastructure, public safety, 178 and shopping options (Deng et al., 2016). With the improvement of urban and inter-179 regional transportation infrastructure in China, regional traffic accessibility is also 180 found to be positively associated with the distribution of talent (Gu et al., 2019a).

181 Recently, scholars have compared the impact of economic opportunities versus amenities on the mobility patterns of talent in China. The cumulative evidence thus far 182 183 suggests that economic opportunities tend to prevail over amenity-driven migration (Gu et al., 2019a; Liu & Shen, 2014b; Yu et al., 2019). However, most prior empirical 184 185 analyses have adopted a cross-sectional perspective on the distribution of talent, neglecting the temporal dynamics of this process. China has continued to experience 186 sustained economic growth in the last decade with focused investment in eastern and 187 188 south-eastern areas of the country (Breznitz & Murphree, 2011). These concentrated 189 poles of economic development are likely to have enticed large flows of talent, 190 generating losses elsewhere, Additionally, recently released population census and 191 sample survey data have rarely been used to investigate the spatial patterns and 192 determinants of talent in China over the last decade. The resources needed to obtain 193 city-level Chinese talent data are costly since the data can only be derived from 194 disaggregated census or 1% sample survey data books of provinces, usually published 195 years after the census year, or the confidential 0.1% micro-level datasets, which are 196 under highly restricted access.

197

198 **2.3 Application of spatial modeling approach in population studies**

199 Spatial autocorrelation is an inherent property in human migration and 200 geographic data. Prior work has consistently shown the presence of strong and positive 201 spatial autocorrelation in the spatial distribution of talent in China (Gu et al., 2019a). 202 Six potential mechanisms that may lead to patterns of spatial autocorrelation have been 203 identified. First, the spatial concentration of talent in an area has been found to lead to 204 a reduction in communication cost of talent in neighboring regions (Fujita & Thisse, 205 1996). Second, it may also lead to an increase in talent volumes in neighboring regions, 206 to take advantage of the existence of agglomeration economies (Miguélez et al., 2010). 207 Third, regional economic growth trigger talent surge as a result of a knowledge spillover effect from adjacent cities with a large stock of talent and technology 208 innovation enterprises (Henderson, 2007). Fourth, mobility of talent searching for jobs 209 210 and residence tends to be spatially focused around their current region of residence and neighboring regions (Liu & Shen, 2017). Fifth, factors influencing the location of talent
in a particular region spills over to nearby areas (Miguélez et al., 2010). Sixth, the interregional social network of talent reinforces concentration around local areas (Gu et al.,
2019a). We recognize all these mechanisms producing spatial autocorrelation may
result from the modifiable areal unit problem as often data used for spatial analysis are
based on arbitrary administrative boundaries (MAUP: Chi & Zhu, 2008, Rowe 2017).

217 When modeling the process of the distribution of talent, traditional econometric 218 models assume independent and identically distributed (i.i.d.) in the model residuals. 219 Without controlling the effect of spatially structured random component (i.e., spatial autocorrelation), model residuals may contain significant spatial autocorrelation, thus 220 violating the i.i.d. assumption and leading to endogeneity (Griffith, 2003; Gu et al., 221 222 2019b). Therefore, appropriate methods should be applied to reduce model bias 223 resulting from spatial autocorrelation, such as the autoregressive model (Anselin, 1988) 224 and the Getis filtering approach (Getis & Griffith, 2002). By adding the selected 225 eigenvectors of the spatial weight matrix as the proxies for spatial autocorrelation, 226 eigenvector spatial filtering (ESF) is another approach for tackling spatial 227 autocorrelation in data. Comparing to other approaches for filtering spatial autocorrelation, ESF is easier to implement for various types of models and does not 228 229 change the estimation method of the model. Thus it is more flexible and has fewer 230 restrictions when capturing spatial autocorrelation (Griffith, 2003).

231

232 **3. Research area, methods, and data**

233 **3.1 Areas**

234 Our research area includes 31 provinces in China, excluding Hong Kong, Macao, 235 and Taiwan. City extents comprise our geographic unit of analysis because they 236 typically represent the spatial scale at which policies are formulated for the attraction 237 and retention of talent, and factors influencing talent often vary significantly accross 238 cities (Gu et al., 2019a; Nie & Liu, 2018). Based on individual-level records from 239 population censuses and sample surveys, the stock of talent is aggregated to the city level. A total of 309 cities are chosen for our basic dataset. Cities in China include 240 241 prefecture cities and autonomous cities. When we do regression analysis, we excluded 242 data on autonomous cities because of their poor quality and number of missing entries. 243 We used a total of 233 cities out of 309 units for regressions. China's economic 244 geography is divided into four economic-geography regions: the Eastern, Central, 245 Western, and Northeastern regions (Figure 1).

- 246 <Figure 1 about here>
- 247

248 **3.2 Methods**

Our analysis is divided into two stages. We first used a range of indicators of inequality to measure the extent of spatial concentration of talent before modeling their migration patterns in a negative binomial framework to identify the key underlying

- factors shaping these patterns, and ESP to account for spatial autocorrelation. Next, wedescribed the indicators of inequality used.
- 254 *Concentration Index (CI).* The CI measures the spatial concentration of talent 255 (Yang et al., 2014). This measure can be expressed in terms of the proportion (%) of 256 talent in 1% of a country's land area. It is represented as:

257
$$CI_i = \frac{P_i/P_n \times 100\%}{A_i/A_n \times 100\%} = \frac{P_i/A_i}{P_n/A_n}$$
 (1)

where CI_i denotes the CI of city *i*; P_i is the talent stock in city *i*; A_i is the land area of city *i*; A_n is the total land area of China; and, P_n is China's total stock of talent.

260 *Coefficient of variation (CV).* The CV complements the CI by measuring the degree 261 of differences in the distribution of talent across the individual spatial unit and is 262 calculated as follows (Haining & Haining, 2003):

263
$$C_{\rm v} = \frac{1}{\bar{p}} \sqrt{\frac{\sum_{i=1}^{n} (P_i - \bar{P})^2}{n-1}}$$
 (2)

where \bar{p} denotes the average stock of talent across cities; *n* is the number of cities; and, p_i is the stock of talent in each city.

Lorenz Curve (LC). The LC graphically represents the degree of inequality of the distribution of talent across China's 309 cities. The LC was generated by plotting the percentage of highly educated population in each city against city layout in ascending order by their talent stocks. The mathematic definition of LC comes from Gastwirth (1971).

271 To model migration flows and identify the patterns underpinning their key drivers, 272 we use a negative binomial model (NBM) framework which is appropriate to model 273 count data (Cameron & Trivedi, 2013): our dependent variables are counts of talent i.e. number of people with a college degree or above. A basic regression modeling structure 274 275 is the Poisson model (PM). However, the underlying assumption of the PM is 276 equidispersion i.e. the variance of the dependent variable is assumed to be equal to the 277 mean (Gu et al., 2019b; Liu & Shen, 2017) and this assumption is usually violated in migration analysis, yielding biased estimations (Rowe 2013b). To address this problem, 278 279 the NBM introduces a parameter α that measures overdispersion in the data. For our 280 dependent variables, the variance was significantly larger than the mean, so that we 281 apply an individual-specific fixed effects NBM to model panel data taking the form of:

282
$$\Pr(\mathbf{Y} = y_{it} | \mu_{it}, \alpha) = \frac{\Gamma(\alpha^{-1} + y_{it})}{\Gamma(\alpha^{-1}) + \Gamma(y_{it} + 1)} (\frac{\alpha^{-1}}{\alpha^{-1} + \mu_{it}})^{\alpha^{-1}} (\frac{\mu_{it}}{\alpha^{-1} + \mu_{it}})^{y_{it}}$$
(3)

283
$$\mu_{it} = \exp\left(\alpha_0 + \beta x_{it} + \varepsilon_i\right) \tag{4}$$

where Γ is the Gamma integral; μ_{it} equals $E(y_{it})$; α is the variance parameter of the Gamma distribution. When this parameter tends to 0, the NBM becomes a PM; y_{it} represents the talent stock of city *i* in time *t*; x_{it} represents a vector of independent variables of city *i* in time *t*; α_0 is the constant term; β represents the vector of estimates capturing the relationship between the independent and dependent variables; and ε_i represents the error term. 290 To correct for spatial autocorrelation in our regression estimates, we use 291 Eigenvector spatial filtering (ESF). The existence of spatial autocorrelation may lead 292 to biased regression estimates, drawing erroneous conclusions and as a result in 293 population migration and distribution data misguided policy recommendations 294 (Griffith, 2003; Casado-Díaz et al. 2017). Eigenvectors of a given spatial weight matrix 295 can represent components with different degrees of spatial autocorrelation. By adding 296 these eigenvectors into regression models, the effect of spatially structured random 297 components is controlled. Hence, ESF has a strong capacity to reduce spatial 298 autocorrelation in residuals (Chun & Griffith, 2011; Fischer & Griffith, 2008; Griffith, 299 2003; Gu et al., 2019b; Liu & Shen, 2017).

In this study, ESF is implemented as follows: first, we constructed an n-by-n binary spatial weight matrix W under *queen* criteria¹ using GeoDa to represent the connectivity of Chinese cities (n denotes the total number of cites). The second step is to construct a transformed spatial weight matrix. Given a matrix $\mathbf{M} = (\mathbf{I} - AA'/n)$, where \mathbf{I} is an n-by-n identity matrix, and A is an n-by-1 vector of 1s. \mathbf{M} can center the spatial weight matrix W by **MWM**. Third, we calculated the eigenvalues and eigenvectors of the matrix **MWM**:

307 $MWM = E\Lambda E'$

(5)

308 where E is the matrix of eigenvectors decomposed by the transformed matrix, and Λ 309 is a diagonal matrix with corresponding eigenvalues.

310 Getis and Griffith (2002) demonstrated that the eigenvectors in the ESF method 311 are orthogonal and uncorrelated, and each eigenvector corresponds to a unique pattern 312 of spatial autocorrelation. Additionally, each individual eigenvector in E can be linked 313 to a Moran's *I*:

314
$$I_j = \frac{n}{A'WA} \frac{e'_j MWMe_j}{e'_j Me_j}$$
(6)

where n is the number of cities. I_i is the Moran's I for the j^{th} eigenvector e_i . The 315 316 eigenvector (e_1) with the largest eigenvalue has the largest Moran's $I(I_1)$ identifying the highest level of spatial autocorrelation in the dataset. The eigenvector (e_2) with the 317 second largest eigenvalue has the second-largest Moran's $I(I_2)$ and indicates the 318 319 second-highest spatial autocorrelation, and so on. In order to choose an appropriate set 320 of eigenvectors controlling spatial autocorrelation in data, Griffith (2003) outlined that 321 candidate eigenvectors can be identified with a critical value of the corresponding 322 eigenvalues indicating a specific minimum spatial autocorrelation level (e.g., Moran's 323 I = 0.25). After the selection of the candidate eigenvectors, a subset of eigenvectors can be chosen with the Akaike information criterion (AIC), and a linear combination of the 324 325 selected eigenvectors can be used to capture latent spatial autocorrelation effects (Chun, 326 2011). This process minimizes the estimation error without reducing too much the degree of freedom of the model. 327

¹ The *queen* spatial weight mareix has been widely applied in defining the spatial relationship of China's cities (Gu et al., 2019a; Hong & Sun, 2011; Liu et al., 2019).

- 328 For space-time panel data analysis (e.g., the NBPM), eigenvectors need to be 329 concatenated T times to match the total number of space-time observations (Chun, 330 2011). Because the spatial structure of the data is invariant over time, linear mixed models (LMM) and generalized linear mixed models (GLMM) are suggested to apply 331 332 to account for spatial and temporal autocorrelation with ESF technique (Chun, 2011; 333 2014; Patuelli et al., 2011). Fortunately, the fixed effects NBPM specification proposed 334 by Hausman et al. (1983) can accommodate the time-invariant spatial autocorrelation 335 represented by the selected eigenvectors (Gu et al., 2019b). After adding eigenvectors that are used to control the effect of spatial autocorrelation, our model specification can 336 337 be defined by:
- 338 $\mu_{it} = \exp(\beta x_{it} + \gamma e_{it} + \varepsilon_i)$

(7)

- 339 where e_{it} denotes the selected eigenvectors of city *i* in time *t*, and γ is the vector of 340 estimators.
- 341

342 3.3 Data

Following existing studies (Gu et al., 2019a; Liu & Shen, 2014a; 2014b), we define talent: individuals with a college degree or above. The rationale of using academic qualifications to define talent is that academic qualifications provide an accurate representation of human capital embedded in the labor force (Schultz, 1961). In most cases, regional development policies of cities often use academic qualifications to define and divide talent (Rowe et al., 2013; Faggian et al. 2016).

We used data from the 2000 and 2010 Chinese population censuses and the 2005 and 2015 1% population sample surveys. In 2000, there were 45.6 million highly educated individuals in China, comprising 3.67% of the total population. By 2005, this number had increased by 30 million to comprise 75.7 million, and by 74 million to represent 120.1 million in 2010. In 2015, the pool of Chinese talent comprised 151.1 million and accounted for 10.99% of the total population.

355 For the independent variables, we used to explain the spatial distribution of talent in China. We draw on data from 1999, 2004, 2009, and 2014 reported in 2000, 2005, 356 357 2010, and 2015 China City Statistical Yearbook (National Bureau of Statistics, 2001, 358 2006, 2011, 2016). Including time-lagged independent variables in our models helps 359 reducing potential endogeneity caused by reverse causation. Additionally, the use of 360 ESF effectively reduces the effect of spatial autocorrelation in residuals and further 361 alleviates any potential endogeneity (Getis & Griffith, 2002; Gu et al., 2019b). As in 362 previous studies (Gu et al., 2019a; Liu & Shen, 2014a; 2014b), we included independent variables to capture two broad sets of factors: economic and amenity factors, and also 363 364 introduced a range of control variables. A description of variables is provided in Table

365 1 for the full set of variables and descriptive statistics are reported in Table 2.

366 <Table 1 about here>

367 <Table 2 about here>

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369 4. The spatial pattern of talent in China

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This section determines the extent of spatial concentration of talent across cities in China. It first delves into the overall characteristics of spatial patterns of talent based on the CI, CV, and LC followed by an analysis of the spatial heterogeneity in these spatial patterns.

374

375 4.1 Concentrated pattern

Following Liu et al. (2010), we divided 309 cities into three types of areas based 376 377 on their CI score: intensively-distributed areas (CI \geq 2), evenly-distributed areas (0.5 \leq $CI \le 2$), and sparsely-distributed areas ($CI \le 0.5$). We analyzed the distribution of talent 378 379 across these areas and the results are reported in Table 3. We found that talent tended 380 to concentrate on a small number of intensively-distributed areas with an increasing 381 density from 32.28 persons/km² in 2000 to 99.97 in 2015. Just over 10% of China's 382 land area clustered about 70% of the highly educated population during the 2000-2015 383 period.

384 <Table 3 about here>

Most of the highly educated population has tended to concentrate in large urban agglomerations and provincial capitals, particularly in the eastern coastal region of the country (Figure 2). This region encompasses intensively-distributed, economically developed areas, including first-tier cities such as Beijing and Shanghai, capital cities in the central and eastern regions such as Zhengzhou, Taiyuan, Wuhan, and Nanjing, and the cities in Pearl River Delta urban agglomerations such as Dongguan, Zhuhai, and Zhongshan.

392 <Figure 2 about here>

Evenly-distributed areas are mainly located in the central-eastern areas of China and have seen a fluctuation in the share of highly educated population, with a slight rise in 2015. Yet, these areas displayed a significant increase in the density of talent . Sparsely-distributed areas emerged persistently in western and north-west parts of the country. They occupy the majority of the land area (more than 60%) but has consistently contained only less than 6% of China's highly educated population. Though they recorded a small increase in density reflecting a reduction in land area.

400

401 **4.2 Unbalanced distribution**

402 We used the CV and LC to measure the degree of spatial inequality of talent. As 403 shown in Figure 3, the CV returned a score of 1.576 in 2000 to 1.634 in 2010 and down 404 to 1.472 in 2015, indicating a persistent pattern of high spatial concentration in the 405 density of highly educated population across cities, with little variation over time. The 406 increasing uneven trend in the first ten years is closely related to a significant difference 407 in college enrollment expansion in particular cities, including first-tier cities (e.g., 408 Beijing and Shanghai) and provincial capitals (e.g., Wuhan and Chengdu) after the introduction of the proposed Action Plan for Education Revitalization in the 21st 409 410 Century. This plan has resulted in a variation in the local supply of educated people 411 across cities. The slight decrease in the concentration of talent between 2010 and 2015

412 may reflect recent national initiatives of regional development to reduce spatial 413 inequalities (e.g., *the National Medium- and Long-Term Talent Development Plan*

414 (2010-2020) and The National New-Type Urbanization Plan (2014-2020).

415 <Figure 3 about here>

416 Similarly, the LC also evidences an unbalanced distribution of talent in China 417 (Figure 4). The vertical axis represents the cumulative percentage of talent stock, while 418 the horizontal axis represents the cumulative percentage of cities ordered by talent stock 419 in each city. The Equality Line (EL) shows a hypothetical balanced distribution of 420 talent. The closer the distance between the LC and EL indicates an evener distribution 421 of talent. The results show a closer LC to the EL for 2015 than for earlier years 422 indicating that a slightly more even distribution of talent in 2015. Despite this decrease, 423 the pattern of talent concentration is still strongly unbalanced and spatially 424 concentrated: a large proportion of the talent (40-45%) concentrated in a small number 425 of large cities (10%).

426 <Figure 4 about here>

427

428 **4.3 Regional heterogeneity**

Underpinning this national pattern of spatial concentration, stark regional differences in the distribution of talent exists. To explore this, we calculated the average density of talent in China's four economic-geography regions from 2000 to 2015 (Figure 5). The eastern region consistently recorded the highest density of talent, displaying a significant increase from 20 persons/km² in 2000 to over 70 in 2015. The western region reports the lowest and most marginal density level remaining below 10 people/km² since 2000.

436 <Figure 5 about here>

437

438 **5. Determinants of talent distribution**

This section explores the determinants of the spatial pattern of talent in Chinese
cities based on the ESF-NPBM, considering the effect of spatial autocorrelation in the
data.

442

443 **5.1 Model processing**

444 The dependent variable is the talent stock in each city (COLLEGE). Drawing on 445 previous studies (Clark & Cosgrove, 1991; Clark & Hunter, 1992; Fielding, 1989; 446 Palivos & Wang, 1996; Partridge, 2010), we defined 14 independent variables. Four variables were used to capture differences in economic opportunities across cities: gross 447 448 domestic product (GDP), average annual wage (WAGE), unemployment rate 449 (UNEMP) and industrial structure (INDUS). Four variables were used to capture 450 differences in public sector amenities: including the ratio of total expenditure on science, technology and education (STEEXPEND), the ratio of fiscal expenditure to 451 452 revenue (SPEND), the number of primary school teachers per 10,000 primary school 453 students (PRIEDU) and the number of doctors per 10,000 people (MEDICAL). We also

454 included three variables to account for differences in natural amenities: green coverage 455 rate (GERRN), sewage treatment compliance rate (SEWAGE) and sulfur dioxide 456 emissions (SO2). Additionally, we introduced three other control variables which may help to explain the distribution of talent: fixed-asset investment (FAI), population 457 458 density (DENS) and the number of college students (UNISTU). To test the robustness 459 of our models, we tested two dependent variables: the number of people with a college 460 degree or above (COLLEGE), and the number of people with an undergraduate degree 461 or above (UNDERGRA).

We tested for strict multi-collinearity of models with the VIF test under the assumption of ordinary least squares (OLS). The VIF values of each variable were all less than 4, indicating that there was no strict multi-collinearity problem. We also calculated the covariance matrix for all variables and found that the pair-wise correlation coefficients did not exceed 0.7. Considering the possible heteroscedasticity, we used a bootstrap technique (resampling 400 times) to compute cluster-robust standard errors (Chen, 2010).

469 To test for spatial autocorrelation, we calculated a panel-type standardized Moran's 470 I for talent stock in 233 cities during 2000-2015 (Arbia & Piras, 2005). The Moran's I for this test was 0.125 (p<0.01), indicating statistically significant and positive spatial 471 autocorrelation in the spatial pattern of talent in cities. To reduce the effect of the spatial 472 autocorrelation, we employed the ESF specification to perform our regression analysis. 473 474 The model results are reported in Table 4. Model 1 corresponds to a standard NBPM, 475 including all 14 explanatory variables but no eigenvector, while for Model 2-6 we used 476 a specification based on ESF and varying groupings of independent variables. The AIC 477 is used to assess model fit. Comparing the AICs for Model 1 and Model 5 returns a 478 much lower AIC for the latter indicating better model fit while controlling for spatial 479 autocorrelation. We thus based our subsequent analyses on the results from the ESF 480 NBPMs. Specifically, we analyzed: Model 2 including only economic-related 481 variables; Model 3 including only amenity-related variables; Model 4 incorporating 482 economic-related and amenity-related variables; Model 5 incorporating control 483 variables; Model 6 based on a different dependent variable (i.e., COLLEGE).

484

485 **5.2 Model results**

486 Results from Model 2 showed statistically significant coefficients for GDP (GDP), 487 average annual wage (WAGE), and unemployment rate (UNEMP). Results from Model 488 3 reveal significant coefficients for scientific technology and education expenditure 489 (STEEXPEND), the number of primary school teachers (PRIEDU), and the number of 490 doctors (MEDICAL). The greening rate (GREEN) and sewage treatment rate 491 (SEWAGE) appear, however, insignificant, while sulfur dioxide emissions (SO2) 492 shows a positive significant correlation to the local count of talent running against our expectation. This finding might be because when the effects of economic opportunities 493 494 were not controlled, sulfur dioxide emissions may have become a good proxy for the 495 local economic level reflecting the correlation between air quality and economic

activity. Results from Model 4 show similar magnitude of coefficients for economicrelated variables as in Model 2, while we observe changes for amenity-related factors,
particularly MEDICAL and SO2 becoming statistically insignificant, and GREEN
resulting in a significantly positive correlation.

500 <Table 4 about here>

501 Comparing Model 5 with Models 2-4 reveals that economic opportunities were 502 the primary factors influencing the spatial distribution of talent in China between 2000 and 2015. GDP reflects higher economic benefits for local communities and population 503 504 in terms of income and consumption spillover effects, generating further consumption in the local economy as well as job creation. The results indicate that an average of a 505 1% increase in urban GDP led to a 0.3202% increase in the stock of talent between 506 507 2000 and 2015. WAGE also exerted a significant impact on shaping the distribution of 508 talent in China, with a 1% increase in the average annual wage of urban employees resulting in a rise in the local talent stock of 0.1848%. Unemployment also seems to 509 510 have played a role. Selecting a workplace is often a risk-averse decision-making process 511 for educated workers; thus our results indicate a negative correlation between the urban 512 unemployment rate and the stock of talent. For every 1% increase in the unemployment rate, the stock of talent reduced by 0.0097%. The industrial structure of the local 513 economy does not seem to play an important role in influencing the pool of local talent, 514 which is partly due to the differences in industrial structure are small across most cities. 515

516 Specific urban amenities were also significantly associated with the spatial 517 distribution of talent. The ratio of per capita science, technology, and education 518 expenditure to financial expenditure (STEEXPEND) emerged as a key factor 519 representing the importance placed by the local government on urban technology and education development. Our results indicate that a 1% increase in this ratio resulted in 520 521 a rise in the stock of talent of the city of 0.0136%. In contemporary China, a large 522 percentage of talent tends to migrate with their family and a primary concern is an 523 education for their children (Gu et al., 2019c). The primary education context of 524 destination cities is thus important in attracting talent. Our results show that if the 525 number of primary school teachers per 10,000 students (PRIEDU) increased by 1%, the 526 pool of talent expands by 0.2556%.

527 Other urban amenities also played a role but to a lesser extent. The rate of urban 528 greening (GREEN) displays a positive relationship with talent stock. Model 5 suggests 529 that a 1% increase in urban greening rate led to a 0.0023% increase in talent stock of 530 the city. However, out of our expectation, other amenity variables were insignificant in 531 our model, including the proportion of per capita fiscal expenditure to fiscal revenue (SPEND), sulfur dioxide emissions (SO2), the number of doctors per 10,000 532 533 (MEDICAL) and sewage treatment rate (SEWAGE), which implied that the 534 relationship between urban amenities and the distribution of talent should be further studied. Particularly, after considering the effects of economic opportunities, GREEN 535 536 becomes significant for model 4 and model 5. Also, the coefficients of SO2 and 537 SEWAGE become positive, although they are insignificant. This indicates that economic opportunities are prerequisites for the influences of natural amenities.
Besides, the effect of MEDICAL becomes insignificant after controlling economicrelated variables, which is partly because the number of doctors per 10,000 is closely
related to a regional economic level.

542

543 5.3 Robustness Checks

We observed the changes in the significance of individual covariates from Model 2 to Model 5 and found that a persistent pattern in the statistical significance and magnitude of economic variables. However, much more variability is observed for amenity-related variables. For instance, coefficients for sulfur dioxide emissions (SO2) and the number of doctors (MEDICAL) become insignificant after controlling for urban economic opportunities.

550 To test the robustness of these results, we used a different proxy for talent i.e. the number of people with a bachelor's degree or above (UNDERGRA) as a dependent 551 552 variable and used the same specification as Model 5. Model 6 in Table 4 reports the 553 results. Consistent with previous models, they show significant coefficients and similar signs for economic factors: GDP, WAGE and UNEMP. Changes are observed for 554 amenity-related variables: GREEN is no longer significant, yet SEWAGE and SO2 555 become significant. This reflects the slight difference between the two types of talent's 556 preferences for amenities. The urban greening rate had a closer relationship with the 557 558 stock of talent with a college degree or above, while the sewage treatment compliance 559 rate and sulfur dioxide emissions were more associated with the stock of talent with a 560 bachelor's degree or above.

561 Overall, based on the results of the ESF NBPMs, urban economic opportunities 562 appear to be essential forces shaping the spatial distribution of talent across Chinese 563 cities. The evidence reveals that only certain amenity-related factors can be consistently 564 be associated with the location of talent. Amenities related to the local provision of 565 primary education, and investment in science, technology, and education seem to be 566 critical.

567

568 6. Conclusion and Discussion

569 This paper assesses and seeks to identify key factors shaping the spatial 570 distribution of talent across Chinese cities between 2000 and 2015. Results revealed a 571 high degree of spatial concentration in the distribution of Chinese talent throughout the 572 fifteen-year period in a small number of cities. While this concentration decreased between 2010 and 2015, it remains relatively high. Examining the density of talent 573 574 within each of four economic-geography regions of China reveals a pattern of rapidly 575 increasing concentration of talent in the eastern region, increasing the gap in the local 576 stock of spatial talent between regions. Only small changes in the local pool of talent were observed in western and north-western cities. 577

578 A significant, positive spatial autocorrelation in the spatial distribution of talent 579 across Chinese cities exists. An ESF specification was used to mitigate the effects of 580 spatial autocorrelation and reduce estimation bias. Our modeling results revealed that urban economic opportunities have played a dominant role in shaping the spatial 581 582 distribution of talent in China. Economic variables were consistently and significantly associated with the local pool of talent. Specifically, regional GDP and average annual 583 584 wage of urban employees showed strong positive relationships with talent stock, while 585 the urban unemployment rate was negatively correlated. This finding is consistent with the evidence from previous cross-sectional studies examining on the spatial distribution 586 587 of China's skilled workers or highly educated people and their determinants, largely 588 prior to 2010 (e.g., Liu & Shen, 2014a). Our work expands previous work by revealing 589 that the systematic key role of economic factors in shaping migration patterns of talent 590 in China over time.

591 Urban amenities seem to have played a less important role. Only specific variables 592 consistently displayed statistically significant associations. Urban per capita science, 593 technology and education expenditure ratio and the number of teachers per 10,000 594 primary school students showed a persistent relationship with the stock of talent in 595 Chinses cities.

596 As China's economic resources and public service supplies are concentrated in 597 developed metropolitan areas with large populations, the existing pattern of spatial concentration and unbalanced distribution in the pool of talent is expected to persist 598 599 over time. Our results have wide-ranging policy implications. From a national and 600 regional perspective, they emphasize the need for appropriate regional development 601 policies to optimize the pattern of talent. For areas with a high density of talent, policies 602 should be targeted at taking advantage of the positive externalities brought about by the 603 clustering of talent. For the central and western regions, where talent is sparsely 604 distributed, preferential policies should be formulated to give specific financial 605 supports to people, such as young graduates to guarantee their fundamental lives. From a city's perspective, local governments should enhance the economic level of cities by 606 offering higher wages and enhancing employment stability to attract talent. Meanwhile, 607 608 governments should also focus on improving public services and the urban 609 environment, particularly primary education amenities and raising local expenditure on 610 science, technology and education, so that the city can meet the increasing and diverse 611 needs of talent.

Using the data of Chinese highly educated talent at the city level, the present paper measures the average effects of economic- and amenity-related factors on the spatial distribution of talent over 2000-2015. China has gone through significant social changes over this period, which may have resulted in a varying effect of the factors shaping the geographic distribution of talent over time. Future studies may estimate the time-variant influence of these factors and how they vary across cities.

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619 **References**

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- 789

	Talent	Land proportion	Density	Density per year	
	proportion (%)	(%)	(persons/km ²)	(persons/km ²)	
2000					
Intensively-distributed area	66.96%	11.60%	32.38		
Evenly-distributed area	28.03%	26.36%	5.97	5.61	
Sparsely-distributed area	5.01%	62.03%	0.45		
2005					
Intensively-distributed area	66.59%	12.09%	46.03		
Evenly-distributed area	27.16%	25.94%	8.75	8.36	
Sparsely-distributed area	6.25%	61.97%	0.84		
2010					
Intensively-distributed area	70.65%	12.30%	85.09		
Evenly-distributed area	23.40%	23.64%	14.67	14.82	
Sparsely-distributed area	5.28%	5.28% 64.06%			
2015					
Intensively-distributed area	65.16%	11.67%	99.97		
Evenly-distributed area	28.84%	27.94%	18.48	17.90	
Sparsely-distributed area	6.00%	60.39%	1.78		

790 Table 1. Concentration pattern of talent from 2000 to 2015

791 792

Table 2. Description and expected effects of variables

Variable	Description	Expected effect
Dependent varia	ables	
COLLEGE	Number of people with a college degree or above of each city in 2000,	
	2005, 2010, 2015	
UNDERGRA	Number of people with a bachelor degree and above of each city in	
	2000, 2005, 2010, 2015	
Economic oppo	rtunity variables	
GDP	Gross GDP of each city in 1999, 2004, 2009 and 2014 (10,000 yuan)	+
WAGE	The average annual wage of employees in urban areas of each city in	+
	1999, 2004, 2009 and 2014 (yuan)	
UNEMP	The urban unemployment rate of each city in 1999, 2004, 2009 and	-
	2014 (%)	
INDUS	The proportion of tertiary industry to GDP of each city in 1999, 2004,	+
	2009 and 2014 (%)	
Amenity variab	les	
STEEXPEND	The proportion of per capita science, technology and education	+
	expenditure to the financial expenditure of each city in 1999, 2004,	
	2009 and 2014 (%)	
CDEND	The ratio of per capita financial expenditure to per capita fiscal revenue	+
SPEND	of each city in 1999, 2004, 2009 and 2014 (%)	

PRIEDU	Number of primary school teachers per 10,000 primary school students	+
PRIEDU	of each city in 1999, 2004, 2009 and 2014	
GREEN	Greening rate of each city in 1999, 2004, 2009 and 2014 (%)	+
CENT CE	Sewage treatment compliance rate of each city in 1999, 2004, 2009	+
SEWAGE	and 2014 (%)	
~~~	Emissions of industrial sulfur dioxide of each city in 1999, 2004, 2009	-
SO2	and 2014 (tons)	
	Number of doctors per 10,000 people of each city in 1999, 2004, 2009	+
MEDICAL	and 2014	
Control variables	8	
DENG	Population density of each city in 1999, 2004, 2009 and 2014	+
DENS	(persons/km ² )	
	Number of college students per 10,000 people of each city in 1999,	+
UNISTU	2004, 2009, and 2014	
<b>D</b> 4 <b>I</b>	Per capita fixed assets investment of each city in 1999, 2004, 2009 and	NS
FAI	2014 (10,000 yuan)	

793 Note: "+" denotes a positive expected effect, "-" denotes a negative expected effect, "NS" represents an unsure
794 expected effect

795

## 796 Table 3. Descriptive statistics of variables

Variable	Number	Mean	Standard	Minimum	Maximum
			deviation		
Dependent varia	bles				
COLLEGE	932	359157	578240.7	11377	7241471
UNDERGRA	932	158344.1	413589.2 2018		6716225
Economic variab	oles				
GDP*	932	12817129.2	21305227.4	168848	235677000
WAGE*	932	24754.39	17273.35	4189	103400.4
UNEMP	928	4.03	2.57	0.05	31.58
INDUS	931	36.49	8.39	12.1	77.95
Amenity variable	es				
STEEXPEND	932	18.13	5.32	2.53	37.36
SPEND	932	213.62	125.06	60.44	1802.50
PRIEDU *	931	559.47	148.53	189.93	1486.94
GREEN	929	35.27	13.90	0.7	92.87
SEWAGE	917	84.32	17.69	0.6	100
SO2*	924	47480.04	27879.18	0.1	641088
MEDICAL *	932	18.65	9.48	2.76	87.51
Control variable	S				
DENS	916	137.95	207.72	0	1293.38
UNISTU*	932	946.23	855.04	12	6161
FAI*	931	17494.63	20967.58	180.61	153802.5

798

## 799 Table 4. Results from the NBPMs and ESF NBPMs

	(1)	(2)	(3)	(4)	(5)	(6)
	COLLEGE	COLLEGE	COLLEGE	COLLEGE	COLLEGE	UNDERGRA
GDP	0.3509***	0.2769***		0.3291***	0.3202***	0.1108*
	(0.0460)	(0.0397)		(0.0458)	(0.0510)	(0.0629)
WAGE	0.1611**	0.3191***		0.1986***	0.1848***	0.6318****
	(0.0643)	(0.0411)		(0.0633)	(0.0695)	(0.0856)
UNEMP	-0.0101*	-0.0114**		-0.0095*	-0.0097**	-0.0194**
	(0.0055)	(0.0051)		(0.0054)	(0.0030)	(0.0090)
INDUS	0.0022	0.0018		0.0015	0.0021	-0.0020
INDUS	(0.0026)	(0.0020)		(0.0032)	(0.0030)	(0.0030)
OTEEVDEND	0.0140***		0.0204**	0.0136***	0.0136***	0.0141***
STEEXPEND	(0.0028)		(0.0046)	(0.0029)	(0.0032)	(0.0056)
CDEND	0.0002		0.0002	0.0002	0.0001	0.0007***
SPEND	(0.0002)		(0.0002)	(0.0002)	(0.002)	(0.0002)
	0.2800***		0.4699***	0.2573***	0.2556***	0.4307***
PRIEDU	(0.0742)		(0.1402)	(0.0753)	(0.0798)	(0.1055)
CDEEN	0.0024**		0.0034	0.0024*	0.0023*	0.0005
GREEN	(0.0011)		(0.0060)	(0.0013)	(0.0011)	(0.0011)
CEWA CE	-0.0006		-0.0005	0.0001	0.0001	0.0026**
SEWAGE	(0.0007)		(0.0010)	(0.0006)	(0.0008)	(0.0011)
50 <b>0</b>	-0.0015		0.0558***	-0.0049	-0.0052	-0.0324***
SO2	(0.0054)		(0.0047)	(0.0044)	(0.0048)	(0.0117)
	-0.0015		0.8173***	-0.0013	-0.0168	0.0533
MEDICAL	(0.0422)		(0.0724)	(0.0543)	(0.0633)	(0.0797)
UNISTU	0.0122**				0.0142**	0.0291***
	(0.0052)				(0.0062)	(0.0099)
E A I	-0.0081				0.0016	-0.0293
FAI	(0.0249)				(0.0301)	(0.0440)
DENG	0.0385*				0.0272	0.1339***
DENS	(0.0213)				(0.0221)	(0.0415)
	-6.5960***	-4.8593***	-4.2884***	-6.3053***	-6.2126***	- 9.8892***
CONSTANT	(0.5038)	(0.2953)	(0.8390)	(0.4238)	(0. 6132)	(0. 7632)
N	892	927	913	909	892	892
Eigenvectors	no	yes	yes	yes	yes	yes
AIC	15903.5	16762.9	16900.38	16274.3	15885.85	14988.07
Log likelihood	-7936.7494	-8361.4506	-8427.1896	-8110.1501	-7910.4024	-7464.0368

800 Note: *** represents p < 0.01, ** represents p < 0.05, * represents p < 0.1; Cluster-robust standard errors are in

801 parentheses.



803 Figure 1. Economic regions in China

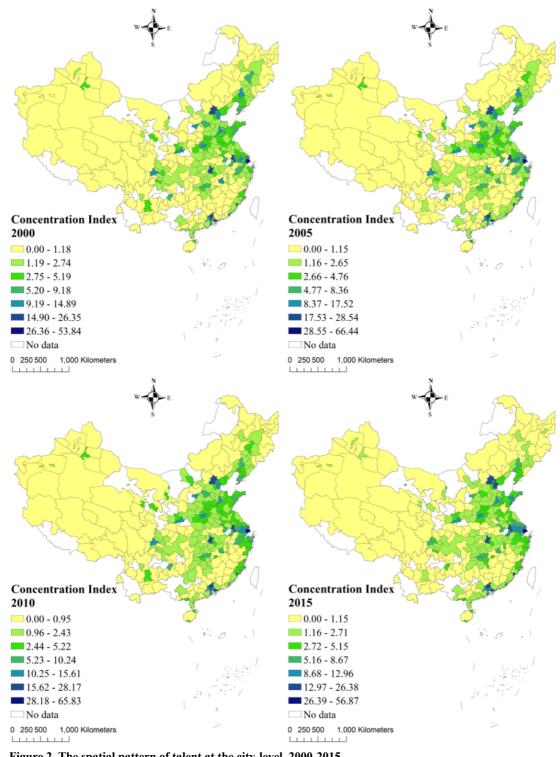




Figure 2. The spatial pattern of talent at the city-level, 2000-2015

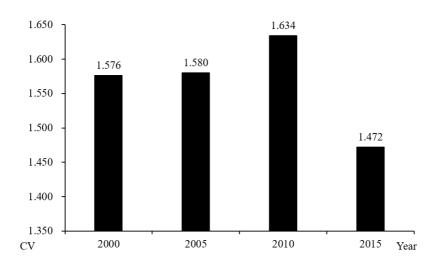
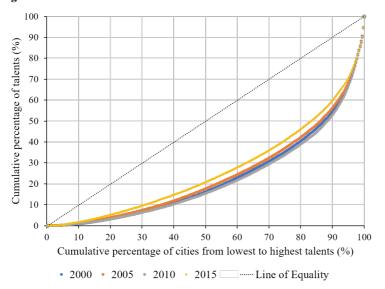




Figure 3. Coefficients of variation of the distribution of talent between 2000 and 2015



809 Figure 4. Lorenz curves of the stock of talent in each city, 2000-2015

