

## Tuning actions and observables in lattice QCD

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We propose a strategy for conducting lattice QCD simulations at fixed volume but variable quark mass so as to investigate the physical effects of dynamical fermions. We present details of techniques which enable this to be carried out effectively, namely the tuning in bare parameter space and efficient stochastic estimation of the fermion determinant. Preliminary results and tests of the method are presented. We discuss further possible applications of these techniques. [S0556-2821(98)10621-5]

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### I. INTRODUCTION

First results from full simulations of lattice QCD have confirmed the magnitude of the computational task ahead and have shown only glimpses of physics beyond the quenched approximation. A recent survey of results can be found in [1].

Preliminary results by the UKQCD Collaboration [2,3] using an  $O(a)$  improved action have shown a surprisingly strong dependence of the effective lattice volume on the bare quark mass. This complicates chiral extrapolations of simulation measurements and obscures comparisons with quenched calculations. With this in mind, we investigate how one might control the effective lattice volume by tuning the bare action parameters while the effects of decreasing quark mass are studied.

Before proceeding, we should clarify what we mean by ‘‘effective lattice volume.’’ For definiteness, consider the Wilson discretization of QCD giving a lattice action dependent on two bare parameters  $\beta$  and  $\kappa$  defined in the usual way. In the quenched approximation, we have become used to thinking of  $\beta$  as uniquely controlling the lattice spacing  $a$ . At fixed  $\beta$ , one makes lattice measurements of the rho mass or the Sommer scale  $r_0$ , defined by [4]

$$r_0^2 \left. \frac{dV}{dr} \right|_{r_0} = 1.65,$$

which is conceptually simpler for the present discussion since, in this case, there is no need to consider extrapolation in the (valence) quark mass. One then matches the lattice value of  $r_0$  to its physical value ( $\approx 0.49$  fm as extracted from heavy quark spectroscopy [4]) to obtain the lattice spacing at that value of  $\beta$ . This mapping  $a_Q(\beta)$  between  $\beta$  and the physical lattice spacing is model dependent in that it is

unique to the quenched approximation. The continuum limit is not accessible directly since the lattice volume vanishes. One must remain at a lattice spacing small enough that discretization errors are small but not so small that finite volume effects are significant.

There are two ways of extending these ideas in the presence of dynamical fermions controlled by the bare mass parameter  $\kappa$ .

(1) The conventional procedure is to construct a similar mapping between  $\beta$  and lattice spacing  $a$  where the matching is made using the lattice value of  $r_0$  extrapolated in  $\kappa$  (sea quark mass) to the chiral limit. This yields a unique, but regularization-dependent, mapping  $a_\chi(\beta)$ . Comparisons with the continuum limit are made as in the quenched case.

(2) Alternatively, one may consider matching the lattice value of  $r_0$  at finite values of  $\kappa$ . Here, the picture is that the simulation is being done with sea quarks of non-infinite mass. Each value of the bare quark mass (or  $\kappa$ ) then corresponds to a different approximation to continuum QCD with light dynamical quarks, in much the same way as does quenched QCD (infinite  $\kappa$ ). Matching in this case results in a mapping  $a(\beta, \kappa)$ . Clearly, this is also regularization-dependent.

The term ‘‘effective lattice volume’’ refers to this second definition of lattice spacing. Our proposal, then, is to conduct simulations in which one attempts to hold the effective lattice spacing fixed. In this way one is better able to keep the physics constant and control lattice artifacts and finite volume effects while studying the effects of light dynamical quarks. In contrast, when adopting the first strategy, the significance of lattice artifacts and finite volume effects changes as the chiral extrapolation is made.

In order to carry out this program one requires a practical way of identifying curves in the  $\beta, \kappa$  plane of constant lattice spacing:

$$a(\beta, \kappa) = \text{const.} \quad (1)$$

Consider  $F$ , the lattice measurement of observable  $f$  with dimension  $d$ , so that

$$F = f a^{-d}. \quad (2)$$

Then, in the scaling region and to leading order, curves of constant  $F$  yield estimates of these curves of constant lattice spacing. In this way, one can track the changes in  $\beta$  required to compensate for changes in  $\kappa$ . As a simple example, one can identify the  $\beta$  shift involved in comparisons of quenched and dynamical simulations. In practice, there will be a residual dependence on the choice of  $F$ . We would expect  $r_0$  to be a ‘‘good’’ choice for exposing a sea quark dependence whereas  $m_\pi$  would *not*, due to the strong dependence on valence quark mass and the effects of chiral symmetry constraints.

In the rest of the paper, we show how the operator and action matching technology introduced in [5] can be used to identify such curves. We demonstrate efficient algorithms for achieving it and present some numerical tests. In Sec. II we summarize the relevant matching formalism required and how it may be used. In Sec. III we describe an efficient algorithm for making stochastic estimates of the fermion determinant [6,7]. Results of numerical tests are presented in the next section. This is followed by a discussion of additional applications of these techniques, including parallel tempering simulations with dynamical fermions. Conclusions and outlook are contained in the final section.

## II. CURVES OF CONSTANT PHYSICS

We first review the matching formalism introduced in [5]. Consider actions  $S_1[U]$  and  $S_2[U]$  describing two lattice gauge theories with the same gauge configuration space  $\{U\}$  so that ( $i=1,2$ )

$$Z_i \equiv \int \mathcal{D}U e^{-S_i[U]}, \quad \langle F \rangle_i \equiv \frac{1}{Z_i} \int \mathcal{D}U F e^{-S_i}. \quad (3)$$

For example,  $S_1$  might be the quenched Wilson action and  $S_2$  the  $O(a)$ -improved action for 2-flavor QCD [9]. In the present application, we will consider  $S_1$  and  $S_2$  to be the same improved fermion action but at different points in the  $\beta, \kappa$  plane. Here,  $F$  is some lattice observable. Expectation values with respect to the two actions can be related via a cumulant expansion whose leading behavior implies [5]

$$\langle F \rangle_2 = \langle F \rangle_1 + \langle \tilde{F} \tilde{\Delta}_{12} \rangle_1 + \dots \quad (4)$$

where

$$\Delta_{12} \equiv S_1 - S_2, \quad \tilde{F} \equiv F - \langle F \rangle, \text{ etc.} \quad (5)$$

In general, an action is a function of several parameters. For example, the Wilson action depends on the bare parameters  $\beta$  and  $\kappa$ . In [5] we considered matching action parameters in one of three distinct ways:

(M1) Match a given set of operators, i.e. require

$$\langle F_n \rangle_1 = \langle F_n \rangle_2.$$

(M2) Minimize the ‘‘distance’’ between the actions, i.e.  $\sigma^2(\Delta_{12})$ .

(M3) Maximize the acceptance in an exact algorithm for  $S_2$  constructed via accept or reject applied to configurations generated with action  $S_1$ .

It was shown that, to lowest order, tuning prescriptions (M2) and (M3) coincide. In fact, if the operators  $F_n$  contribute to the action with weights which are considered as tuning parameters, then prescription (M1) also coincides to lowest order. The prescriptions differ in a calculable way at next order. Details are in [5].

In the present application we take

$$S_1 = S_{\text{eff}}(\beta_0, \kappa_0) \quad \text{and} \quad S_2 = S_{\text{eff}}(\beta, \kappa) \quad (6)$$

and seek to explore the bare parameter dependence of the lattice theory using configurations generated at a series of reference points  $(\beta_0, \kappa_0)$  in parameter space.

Here,  $S_{\text{eff}}$  is the effective action corresponding to lattice QCD with the fermions integrated out: i.e.

$$S_{\text{eff}} = -\beta W_\square - T \quad (7)$$

where  $W_\square$  is the usual Wilson plaquette action,

$$W_\square \equiv \frac{1}{3} \sum_{\square} \text{Re Tr } U_\square, \quad (8)$$

and

$$T \equiv n_f \text{Tr} \ln M[U] = \frac{n_f}{2} \text{Tr} \ln(M^\dagger M). \quad (9)$$

The fermion matrix  $M$  for the non-perturbative  $O(a)$  improved theory is a function of both  $\kappa$  and  $\beta$ . This is because the parameter  $c_{\text{sw}}$  [9] is a function of  $\beta$  [10]. Since the improvement scheme fixes  $c_{\text{sw}}(\beta)$ , one must not treat  $c_{\text{sw}}$  as an independently tunable parameter. However, as we shall see, the fact that the operator  $T(\beta, \kappa)$  is a function of  $\beta$  as well as  $\kappa$  introduces some practical complications.

According to Eq. (4), one requires measurements of

$$\Delta_{12} \equiv S_1 - S_2 = (\beta - \beta_0) W_\square + T(\beta, \kappa) - T(\beta_0, \kappa_0) \quad (10)$$

in order to carry out parameter tuning. We discuss efficient algorithms for this in Sec. III.

For now, consider the matching lattice observable  $F$  at two neighboring points in the  $\beta, \kappa$  plane:

$$(\beta_0, \kappa_0) \quad \text{and} \quad (\beta, \kappa) \equiv (\beta_0 + \delta\beta, \kappa_0 + \delta\kappa). \quad (11)$$

According to prescription (M1) above and Eq. (4) we require, to first order in small quantities,

$$\langle \tilde{F} \tilde{\Delta}_{12} \rangle_1 = 0. \quad (12)$$

From this we can deduce that the constant  $F$  curve is given by

$$\frac{\delta\beta}{\delta\kappa} = - \frac{\langle \tilde{F} \delta\tilde{T} \rangle_1}{\delta\kappa \langle \tilde{F} \tilde{W}_\square \rangle_1} \quad (13)$$

where

$$\delta\tilde{T} = \tilde{T}(\beta_0 + \delta\beta, \kappa_0 + \delta\kappa) - \tilde{T}(\beta_0, \kappa_0). \quad (14)$$

Equation (13) amounts to a non-linear differential equation since the right hand side involves  $\delta\beta$  (via the  $c_{sw}$  parameter). Linearizing and taking the limit yields, for the constant  $F$  curve,

$$\frac{d\beta}{d\kappa} = - \frac{\frac{\partial \langle \tilde{F} \rangle}{\partial \kappa}}{\frac{\partial \langle \tilde{F} \rangle}{\partial \beta}} = - \frac{\left\langle \tilde{F} \frac{\partial \tilde{T}}{\partial \kappa} \right\rangle}{\left\langle \tilde{F} \left( \tilde{W}_\square + \frac{\partial \tilde{T}}{\partial c_{sw}} \dot{c}_{sw} \right) \right\rangle}. \quad (15)$$

The quantity

$$\dot{c}_{sw} = \frac{dc_{sw}}{d\beta}$$

is well determined [10] and so the determination of constant  $F$  curves reduces to measuring correlations of the form

$$\langle \tilde{F} \tilde{W}_\square \rangle_1 \quad \text{and} \quad \langle \tilde{F} \delta\tilde{T} \rangle_1. \quad (16)$$

The details of this will be described in Sec. IV.

As pointed out in the Introduction, the details of these curves will depend on the choice of  $F$ . For sensible choices and reasonably physical values of the parameters, one would hope that the corresponding curves of constant  $a$ ,  $a_F(\beta, \kappa)$  say, would agree rather closely, locally at least. For demonstration purposes, we will consider in Sec. IV several simple choices for  $F$ :

- $P$ , the average plaquette, proportional to  $W_\square$ .
- $W_L$ , various Wilson loops.
- $S_{\text{eff}}$ , the complete effective action itself.
- Correlation matrices yielding the static potential and  $r_0$ .
- Hadron correlators.

The first of these,  $P$ , may be readily measured with high accuracy and so is excellent for testing the basic technology. However, it is not expected to shed much light on lattice spacing. The last two are computationally more demanding but more relevant to the project at hand, identifying curves of fixed physical volume.

One might expect that matching the full action ( $F = S_{\text{eff}}$ ) would be more physically relevant than matching the plaquette piece of the action. From Eq. (15) we see that this curve is determined by correlations of the form

$$\langle \tilde{T} \delta\tilde{T} \rangle_1 \quad (17)$$

in addition to those of (16). As we shall see in Sec. IV, it is more difficult to obtain unbiased estimators for these.

Now consider matching scheme (M2) where ‘‘distances’’ in the action space are minimized. One can think of this as

defining ‘‘geodesics’’ in the  $\beta, \kappa$  plane with respect to the metric implied by Eqs. (3), i.e.

$$g_{\mu\nu} = \langle (\partial_\mu \tilde{S})(\partial_\nu \tilde{S}) \rangle.$$

The corresponding affine connection would be

$$\Gamma_{\mu\alpha\beta} = \langle (\partial_\mu \tilde{S})(\partial_{\alpha\beta} \tilde{S}) \rangle.$$

Simple minimization yields, to first order,

$$\frac{d\beta}{d\kappa} = - \frac{\left\langle \left( \tilde{W}_\square + \frac{\partial \tilde{T}}{\partial c_{sw}} \dot{c}_{sw} \right) \frac{\partial \tilde{T}}{\partial \kappa} \right\rangle}{\left\langle \left( \tilde{W}_\square + \frac{\partial \tilde{T}}{\partial c_{sw}} \dot{c}_{sw} \right)^2 \right\rangle}. \quad (18)$$

In Sec. V we show that these curves are directly relevant to simulations of full QCD using parallel tempering. Again, we note that these curves involve operator correlations which are more complex to estimate.

Finally in this section, we observe that some of the above formalism simplifies considerably in the case of the unimproved Wilson action ( $c_{sw} = 0$ ) since then, for example,

$$\frac{\partial F}{\partial \beta} = \langle \tilde{F} \tilde{W}_\square \rangle. \quad (19)$$

### III. STOCHASTIC ESTIMATOR OF THE FERMION DETERMINANT

We require an unbiased estimator for  $T = \text{Tr Ln } H$  where  $H = M^\dagger M$  is a Hermitian positive-definite matrix. We will also require estimators for  $T^2$ ,  $\delta T$  and  $T \delta T$ . Bai, Fahey and Golub [6] have recently proposed estimators, with bounds, for quantities of the form

$$u^\dagger g(H) v \quad (20)$$

where  $g$  is some matrix function. In our application  $g$  is the logarithm and, for convenience of notation, we set

$$L \equiv \ln(H). \quad (21)$$

Taking  $u = v = \phi_i$ , some normalized noise vector (e.g.  $Z_2$  or Gaussian), we can obtain a stochastic estimate of  $T$  via

$$E_T = \frac{1}{N_\phi} \sum_{i=1}^{N_\phi} \phi_i^\dagger L \phi_i. \quad (22)$$

The corresponding variance of this estimator is

$$\sigma^2(E_T) = \frac{1}{N_\phi} \text{Tr}(L^2) \quad (23)$$

for complex Gaussian noise, and something less than this for  $Z_2$  noise ( $\pm 1$  on each of the complex components). In the case of Gaussian noise, we also obtain rather directly an efficient unbiased estimator for  $T^2$ ,

$$E_{T^2} = (E_T)^2 - \frac{1}{N_\phi} E_Q, \quad (24) \quad \langle \hat{E}_{T^2} \rangle_\phi \approx \langle E_{T^2} \rangle_\phi = T^2. \quad (30)$$

where  $E_Q$  is an unbiased estimator for  $Q = \text{Tr } L^2$ ,

$$E_Q = \frac{1}{N_\phi} \sum_{i=1}^{N_\phi} \phi_i^\dagger L^2 \phi_i. \quad (25)$$

In the case of  $Z_2$  noise, the corresponding estimator is not readily accessible via the techniques described below; so we restrict the discussion to complex Gaussian noise.

In a companion paper [11], we give fuller details of methods for evaluating the quantities  $\phi_i^\dagger L \phi_i$ ,  $\phi_i^\dagger L^2 \phi_i$  and so on. In practice, we make no use of the bounds presented in [6]. Instead we use large enough Lanczos systems so that numerical convergence renders the bounds irrelevant. The efficiency of the method results from an elegant relationship between the nodes and weights required for an  $N$ -point Gaussian quadrature and the eigenvalues and eigenvectors (so-called Ritz pairs) of a Lanczos matrix of dimension  $N$  [6]. In the companion paper [11], we show that this relationship and resulting accuracy remain good even when orthogonality is lost. This is an important point for our application. If the Lanczos system is large enough to avoid truncation errors, one is well into the regime where orthogonality is lost in standard numerical Lanczos methods. In the present paper we merely summarize the formulas required to obtain the present set of results. Preliminary results using these techniques were presented in [7].

The actual estimator which we use for  $T \equiv \text{Tr } L$  is

$$\hat{E}_T = \frac{1}{N_\phi} \sum_{i=1}^{N_\phi} I(\phi_i) \quad (26)$$

where

$$I(\phi_i) = \sum_{j=1}^N \omega_j^2 \ln(\lambda_j). \quad (27)$$

Here  $\{\lambda_j^i\}$  ( $j=1,2,\dots,N$ ) are the eigenvalues of the  $N$ -dimensional tridiagonal Lanczos matrix formed using  $\phi_i$  as a starting vector. The weights  $\{\omega_j^2\}$  are related to the corresponding eigenvectors [6]. In fact,  $\omega_j$  is just the first component of the  $j$ th eigenvalue of the tridiagonal matrix.

The estimator  $\hat{E}_Q$ , for  $Q \equiv \text{Tr } L^2$ , is obtained from Eqs. (26) and (27) using  $\ln(\lambda_j^2)$  rather than  $\ln(\lambda_j)$ . For  $T^2$  [see Eq. (24)], we define

$$\hat{E}_{T^2} = (\hat{E}_T)^2 - \frac{1}{N_\phi} \hat{E}_Q. \quad (28)$$

It is straightforward to show that, with the above definitions,

$$\langle \hat{E}_T \rangle_\phi \approx \langle E_T \rangle_\phi = T \quad (29)$$

and

The above Lanczos-based methods for evaluating  $\phi_i^\dagger L \phi_i$  are significantly more efficient than the Chebychev-based methods used previously [8,5]. For a given level of accuracy in the present applications, they are between 3 and 5 times more economical in the number of matrix multiplications required. Typically we achieve 6 figure convergence of the quadrature with 70 Lanczos steps on a matrix with a condition number ( $\lambda_{\max}/\lambda_{\min}$ ) of order  $10^4$  or  $10^5$ .

Our goal was to achieve variance with respect to  $\phi$  [see Eq. (23)] which was one order of magnitude less than that with respect to the physical (gauge) distribution. We found that  $N_\phi = 80$  was a suitably conservative number of noise vectors to use.

For estimating  $\delta T$ , we note that

$$E_{\delta T} = E_{T'} - E_T = \frac{1}{N_\phi} \sum_{i=1}^{N_\phi} \phi_i^\dagger (L' - L) \phi_i. \quad (31)$$

Thus we can achieve variance

$$\sigma^2(E_{\delta T}) = \frac{1}{N_\phi} \text{Tr}[(L' - L)^2] \quad (32)$$

if we use

$$\hat{E}_{\delta T} = \hat{E}_{T'} - \hat{E}_T \quad (33)$$

provided we have employed the *same* set of noise vectors. This is simple to arrange.

In fact, one could use stochastic estimators of the form (20) to estimate directly

$$\frac{\partial T}{\partial \kappa} = \frac{n_f}{\kappa} \text{Tr}[1 - M^{-1}] = \frac{n_f}{\kappa} \text{Re } \text{Tr}[1 - M(M^\dagger M)^{-1}]. \quad (34)$$

An estimator for  $\text{Tr}[M(M^\dagger M)^{-1}]$  is obtained by setting

$$v = \phi, \quad u = M^\dagger \phi \quad (35)$$

where  $\phi$  is a suitable (e.g. Gaussian) noise vector. Then

$$u^\dagger (M^\dagger M)^{-1} v = v^\dagger M (M^\dagger M)^{-1} v \quad (36)$$

is an unbiased estimator of  $\text{Tr}[M(M^\dagger M)^{-1}]$  as required. For this unsymmetric case ( $u \neq v$ ), two Lanczos systems must be used and a subtraction performed [6]. For the present analysis we have used the symmetric formalism as described above. Further analysis and discussion of these and related stochastic estimators are presented in [11].

In the next section we report results of some tests of the matching procedures using the above algorithms.

TABLE I. Parameters association with HMC simulation and with the stochastic estimation of the determinant.

Parameter	Description	Value
$N_{\text{step}}$	Number of HMC steps per trajectory	50
$N_{\text{solve}}$	Number of sweeps in the HMC solver (e.g. BiCGStab)	300
$A$	HMC trajectory acceptance	0.75
$\tau_{\text{AC}}$	Autocorrelation time in trajectories	30
$N_{\phi}$	Number of noise vectors (Sec. III)	80
$N_{\text{Lanc}}$	Number of Lanczos iterations	90

#### IV. NUMERICAL TESTS OF MATCHING

##### A. Work estimates

Having generated an ensemble of decorrelated configurations, one can consider estimating the numerical derivatives required for matching [see Eq. (15)] in one of two ways:

(1) Conduct two further simulations at neighboring points in parameter space and make (uncorrelated) measurements of  $\langle F \rangle$  on each of these.

(2) Apply the stochastic trace logarithmic techniques of the previous section to the existing ensemble.

One can estimate the relative amount of work involved in these two approaches. Suppose we seek to achieve an absolute error of  $\epsilon$  on a measurement of  $\delta\langle F \rangle$  by each method and that the variances of  $\langle F \rangle$  and  $\langle \delta\tilde{T}\tilde{F} \rangle$  are  $\sigma_F^2$  and  $\sigma_{\delta F}^2$  respectively. The ratio of work required for the two approaches is then

$$\frac{W_2}{W_1} = \frac{n_T}{4} \frac{\sigma_{\delta F}^2}{\sigma_F^2} \frac{W_T}{W_{\text{HMC}}} \quad (37)$$

where  $W_T$  is the work done in a stochastic estimate of the trace logarithm on one configuration,  $n_T$  is the number of trace logarithms required (either 2 or 3) and  $W_{\text{HMC}}$  is the work done in generating one decorrelated configuration by hybrid Monte Carlo (HMC) simulation. In turn, we estimate

$$\frac{W_T}{W_{\text{HMC}}} = \frac{N_{\phi} N_{\text{Lanc}}}{N_{\text{step}} N_{\text{solve}}} \frac{A}{\tau_{\text{AC}}} \quad (38)$$

where the various parameters are associated with HMC simulation and the stochastic estimation of the determinant. These are defined in Table I which also shows typical values from the analysis presented below. With the numbers shown, the ratio  $W_T/W_{\text{HMC}}$  is about 1/80. In Table II we show

TABLE II. Sample variances used to estimate the relative work involved in measuring differences directly and by the techniques proposed in this paper. These correspond to a shift  $\delta\kappa = -.0005$  and the reference data described in the text.

$F$	$\sigma_F^2$	$\sigma_{\delta F}^2$	$\sigma_{\delta F}^2/\sigma_F^2$
$P$ (average plaquette)	$5.2 \times 10^{-6}$	$2.9 \times 10^{-5}$	5.5
$r_0$	0.78	2.7	3.5

sample values of the variances  $\sigma_F^2$  and  $\sigma_{\delta F}^2$  taken from the present study. Putting all this together we estimate

$$\frac{W_2}{W_1} \approx 0.045. \quad (39)$$

Although the variance of  $\langle \delta\tilde{T}\tilde{F} \rangle$  can be quite large on large lattices, it is proportional to  $(\delta\kappa)^2$ , or whatever the relevant difference parameter is. It is therefore possible, in principle, to obtain an acceptable accuracy for the relevant derivative by method (2) with significantly less work. The above example suggests this is less than 5% of that required to obtain comparable accuracy by making additional simulations. One can easily refine the above treatment to take account of the work involved in equilibration. Of course, this makes method (2) seem even more attractive.

##### B. Fixed plaquette curve

First, we present results from matching the average plaquette  $\langle P \rangle$ . This provides a check on some basic features of the procedures: the first order truncation of the cumulant expansion (4) and the Lanczos-based noisy estimator algorithm (26). The initial reference point in the  $\beta, \kappa$  plane was taken as

$$(\beta_0, \kappa_0) = (5.2, 0.1340) \quad (40)$$

where a sequence of well-equilibrated configurations had been generated by hybrid Monte Carlo simulation on a  $8^3 24$  lattice with a non-perturbatively improved action for 2 flavors [3]. At  $\beta=5.2$ , the relevant improvement coefficient is  $c_{\text{sw}}=2.0171$  [10]. It is estimated from preliminary spectroscopy measurements on larger lattices that  $\kappa_{\text{crit}}$  is around 0.1365 at this  $\beta$  and  $c_{\text{sw}}$ . The values of  $\kappa$  considered in this analysis are typical of those being used in production simulations. Plaquette measurements from all trajectories within this data sample showed autocorrelation times less than 20 trajectories. Longer runs showed an autocorrelation time for the plaquette of roughly double this. A bootstrap approximation with binning was used to estimate the errors on all quantities.

Using the algorithm and choices described in Sec. III, we made stochastic estimates of  $T \equiv \text{Tr} \ln M^\dagger M$  on a sequence of 40 configurations separated by 20 trajectories. In order to make use of Eqs. (13) and (15) this must be done using a minimum of  $n_T=3$  parameter sets:

$$\{c_{\text{sw}}(\beta_0), \kappa_0\}, \quad \{c_{\text{sw}}(\beta_0), \kappa_0 + \delta\kappa\}, \quad \{c_{\text{sw}}(\beta_0 + \delta\beta), \kappa_0\}. \quad (41)$$

Recall that for the unimproved Wilson action, there is no need to account for the additional  $\beta$  dependence in  $T$  which enters via  $c_{\text{sw}}(\beta)$ .

Having selected a change in bare quark mass  $\delta\kappa$  (for example  $\delta\kappa = -0.0005$ ), we then use Eq. (15) with  $F \equiv P$  to estimate  $d\beta/d\kappa$ . This yields a first estimate of the change  $\delta\beta$  required to maintain a fixed value of  $\langle P \rangle$ . One can then use Eq. (13) to verify this estimate of  $\delta\beta$  by making further stochastic estimates of  $T$  at the final parameter set  $\{c_{\text{sw}}(\beta_0$

TABLE III. Matching the average plaquette.

$(\beta_0, c_{sw}, \kappa_0)$	$\langle P \rangle$	Trajectory	$\kappa$	$\delta\kappa$	$\delta\beta$	Configuration
(5.200, 2.0171, 0.1340)	0.5286(3)	6000	0.1335	-0.0005	+0.0199(10)	40
			0		+0.41(3)	40
(5.220, 1.9936, 0.1335)	0.5290(3)	6000	0.1340	+0.0005	-0.0195(15)	40
			0.1330	-0.0005	+0.0162(23)	40
			0		+0.40(4)	40
(5.61, 0)	0.5275(3)	1000 sweeps				
(5.200, 2.0171, 0.1330)	0.5197(3)	6000	0.1325	+0.015(3)		40
			0	+0.368(16)		40
(5.215, 1.9994, 0.1325)	0.5207(6)	2000				

+  $\delta\beta$ ,  $\kappa + \delta\kappa$ }. In all cases studied, this last verification step has been well satisfied and so one can in fact identify the matching curve directly from the two partial derivatives as proposed in the previous section. Results are shown in Table III.

We have then generated further dynamical fermion configurations at the matched point

$$(\beta_0 + \delta\beta, \kappa_0 + \delta\kappa) = (5.220, 0.1335) \quad (42)$$

and accumulated a similar ensemble of configurations for subsequent measurement. The corresponding value of  $c_{sw}$  is 1.9936. The plaquette measurements were made using relatively high statistics, yielding the statistical errors shown. In order to make a proper comparison, one should fold into the error on  $\langle P \rangle$  at (5.220, 1.9936, 0.1335) that due to the uncertainty in  $\beta$  ( $\pm 0.0010$ ). This would feed through to an additional uncertainty in  $\langle P \rangle$  of  $\pm 0.0004$ . Thus the matching test is very well satisfied for the plaquette.

The matching prediction done in reverse, back from  $\kappa = .1335$  to .1340, is also seen to be well satisfied. From Table III we see that (5.220, 0.1335) is expected to match with (5.205(15), 0.1340) in good agreement with (5.2, 0.1340) from where the matching estimates were originally made. Also in Table III, we show the estimated  $\beta$  shift corresponding to a further change of  $-0.0005$  in  $\kappa$ . Steps of this kind allow one to set up a grid of points from which one can then deduce the fixed plaquette curve in the  $\beta, \kappa$  plane.

For the sake of completeness, we also show in Table III the estimated shift required to match with quenched measurements of  $\langle P \rangle$  (i.e at infinite  $\kappa$ ). Independent gauge simulations at  $\beta=5.61$  show a good match of  $\langle P \rangle$  when one takes into account the additional uncertainty in  $\langle P \rangle$  of around  $\pm 0.009$  which would feed through from the error of  $\pm 0.03$  in estimating  $\beta$ . Of course, for such large shifts a first order approximation may not be sufficiently accurate. We have calculated the second order approximation to  $\delta\beta$  [5] for this quantity [ $+0.52(10)$ ] but the statistical error is such that one cannot reliably discriminate it from the first order result [ $+0.41(3)$ ] with the present level of statistics.

A further plaquette matching test is given in Table III. In this example, the reference point for the ‘‘constant’’ plaquette curve was (5.2, 0.1330). Again, direct simulation

showed that the matching was accurate and self-consistent. The matched points on this curve [ $\langle P \rangle = 0.5197(3)$ ] have been used to conduct tests of parallel tempering as described in Sec. V.

### C. Full action matching

We have noted in Sec. II that some of the correlations required to identify curves of constant action (18) are not directly calculable by the techniques of Sec. III. Those of the form

$$\langle T(c_{sw}, \kappa) T(c'_{sw}, \kappa') \rangle \quad (43)$$

require some care when setting up unbiased estimators. In particular,

$$E_{TT'} \equiv E_T E_{T'} \quad (44)$$

is *not* unbiased. For an unbiased estimator, one requires something like that used for  $T^2$ , i.e.  $E_{T^2}$ . Unfortunately it is not so easy to evaluate the analogue of Eq. (25) via the above Lanczos methods. However, provided the variance of  $E_T$  with respect to noise (23) is indeed small compared to that with respect to gauge fluctuations, the estimator (44) provides a useful approximation. Using this approximation, we have measured the shift  $\delta\beta$  corresponding to a shift  $\delta\kappa$  at fixed action to compare with that at fixed plaquette. For the first test shown in Table III, we find  $\delta\beta = 0.0200(10)$  (statistical error only), consistent with the value 0.0199(10) found for the fixed plaquette curve. In each case studied, we have found such consistency.

We conclude that the fixed action and fixed plaquette curves are not significantly different at this order.

### D. Gauge invariant loops

We are especially interested in matching those observables which are more sensitive to long range physics. We have repeated the above plaquette matching analysis using a variety of Wilson loops. The 16 loops used consist of 4 basic shapes realized in 4 different ‘‘magnifications’’ ( $\times 1$ ,  $\times 2$ ,  $\times 3$  and  $\times 4$ ). The 4 basic shapes used were those specified in terms of link steps by the operations shown in Table IV and rotations thereof.

TABLE IV. Construction of sample gauge invariant loops from specified link steps. A step +2 means a link along lattice direction 2 while -1 means a link along lattice axis 1 in the negative direction.

Loop	No. of links	Link steps
1	4	(+1,+2,-1,-2)
2	6	(+1,+1,+2,-1,-1,-2)
3	6	(+1,+2,+3,-2,-1,-3)
4	6	(+1,+2,+3,-1,-2,-3)

For each loop  $L$ , we have evaluated the shift  $\delta\beta_L$  required to hold the Wilson loop  $\langle W_L \rangle$  constant under a change  $\delta\kappa$  ( $= -0.0005$ ). The results of this are shown in Fig. 1. The values of  $\delta\beta_L$  are similar to each other and to that for the average plaquette measured above. There is not much evidence of a shift increasing with the loop size, although the  $\times 2$  loops show a higher trend than the plaquette value. One sees little evidence of mis-matching above the level of a standard deviation.

If this result (same  $\delta\beta$ ) was reproduced for *all* gauge invariant loops and *all* linear combinations thereof, we would conclude that the static potential  $r_0$ , and hence the lattice spacing, would be identical at the matched points (40) and (42). This would realize our initial objective of defining curves of constant effective volume. However, it cannot be that *all* fixed- $F$  curves emanating from a finite reference point  $(\beta_0, \kappa_0)$  coincide. Moving from this point, in one direction we approach the quenched limit and in the other the chiral limit. We expect that different observables will be more or less sensitive to the effects of quenching. For example, the mass of a vector meson will probably change by less than 10% as the chiral limit is reached in the full theory as compared to its quenched value. On the other hand, the

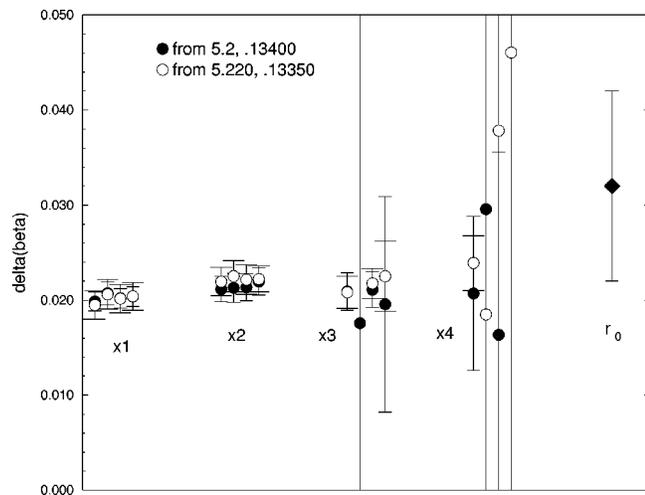


FIG. 1. Predictions for the shift  $\delta\beta$  from reference point (5.2, 0.1340) obtained using the 16 sample loops described in the text. The diamond point is the corresponding value deduced from  $r_0$ . The solid (open) points correspond to a shift of  $\delta\kappa = -0.0005$  ( $+0.0005$ ) from ensembles at (5.2, 0.1340) and (5.220, 0.1335) respectively.

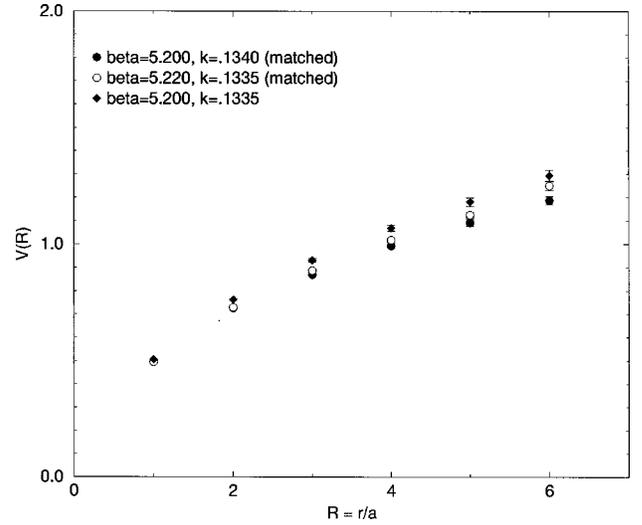


FIG. 2. Static potential on an  $8^3 \times 24$  lattice at (5.2, 0.1340) and (5.220, 0.1335) where the average plaquette values match.

string tension should change from the lattice equivalent of 440 MeV to zero, eventually. We study the static potential in the next section.

### E. Potential and $r_0$

We have used the methods of [12] to measure the potential on each of the main ensembles studied. Since these are on  $8^3 \times 24$  lattices, there are strong finite size effects present in  $V(r)$  and  $r_0$  at the parameter values of interest. However, for the present purpose this is of little consequence. In the variational methods of [12] one constructs “fuzzed” loops from a variety of spatial paths and employs transfer matrix methods to extract energy eigenvalues. The potential values were estimated by taking weighted averages of the effective masses at large time. We took care to use the same procedures on all ensembles. Errors were estimated by the bootstrap method.

Figure 2 shows the static potential at the matched points (40) and (42). The values are in good agreement at short distances but show a systematic divergence at larger separations. Figure 3 shows more clearly the difference between the potentials  $V(5.2, 0.1340) - V(5.220, 0.1335)$ . There appears to be a systematically increasing difference at larger distance. For comparison, the figure also shows  $V(5.2, 0.1340) - V(5.2, 0.1335)$  where there has been a shift in  $\kappa$  but *no* compensating change in  $\beta$ . There is clear disagreement at all distances, as expected.

The remaining set of points in Fig. 3 shows the prediction for (5.220, 0.1335) from the reference ensemble at (5.2, 0.13400) using Eq. (4) to first order. Within the large statistical errors, the predicted difference is indeed compatible with zero. This demonstrates that where matching has been done only approximately (in this case with  $\langle P \rangle$ ) an observable can still be reliably estimated at another nearby point of interest.

From the comparison of the directly simulated points, we conclude that matching the plaquette is not equivalent to matching the long range potential. Note that no corrections

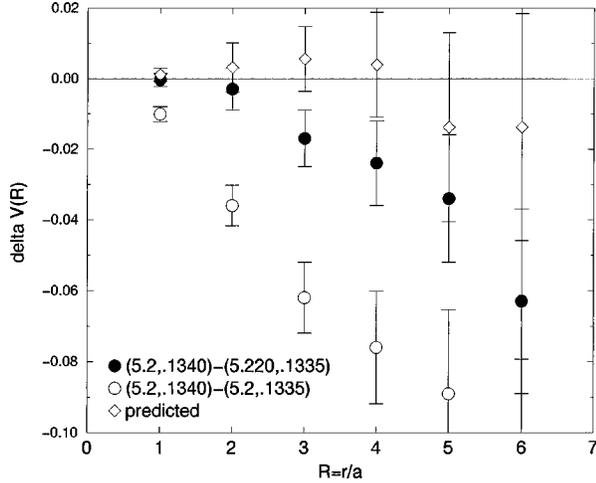


FIG. 3. Differences in the static potential measured from the reference point (5.2, 0.1340). Solid and open circles correspond to (5.220, 0.1335) and (5.2, 0.1335) respectively. Open diamonds correspond to the prediction for (5.220, 0.1335) using Eq. (4).

have been made for lattice artifacts at short distances but these are expected to be similar in each case.

From the potential measurements represented in Fig. 2 we can extract corresponding values of  $r_0$  [4] and, hence, lattice spacing using  $r_0(\text{phys})=0.49$  fm. These are shown in Table V. In finding  $r_0$ , we have calculated interpolated values of the static force via Newtonian 5-point interpolation of the potential. As usual, errors were computed via the bootstrap method. The same procedures were used for each ensemble.

From Table V, we see that the lattice volume at the matched points is similar although that at the heavier quark mass is perhaps one standard deviation larger. This is a reflection of the observation made above that the potentials diverge at large separations due to their different slopes. We have attempted to estimate the  $\delta\beta$  required to match the mean values of the lattice spacings. In the last three rows of Table V, we show the lattice spacing predicted from the reference ensemble via Eq. (4) for  $\beta$  shifts of +0.02, +0.03 and +0.04 to go with the  $\kappa$  shift of  $-0.0005$ . The spacing predicted for  $\delta\beta=0.02$  agrees well with that measured by direct simulation (previous line). From this value and those corresponding to  $\delta\beta=0.03$  and 0.04 we estimate that the

TABLE V. Values of  $r_0$  and lattice spacing deduced from the static potential. The last two rows are values predicted by Eq. (4) from the reference ensemble (5.2, 0.13400).

$(\beta, \kappa)$	$r_0/a$	$a$ (fm)	
(5.200, 0.1340)	3.87(17)	0.127(6)	Direct simulation
(5.220, 0.1335)	3.69(11)	0.133(4)	by HMC
(5.232, 0.1335)	3.76(13)	0.130(4)	
(5.61, 0)	2.33(2)	0.211(2)	
(5.220, 0.1335)	3.67(18)	0.134(6)	Prediction
(5.230, 0.1335)	3.80(22)	0.129(9)	from (5.2, 0.13400)
(5.240, 0.1335)	4.06(33)	0.121(12)	using Eq. (4)

TABLE VI. Summary of the shifts  $\delta\beta$  required to maintain the specified quantities constant under a change  $\delta\kappa=-0.0005$  from the reference point  $(\beta_0, \kappa_0)=(5.200, 0.13400)$ .

Measurement	Value at $(\beta_0, \kappa_0)$	$\delta\beta$
$\langle P \rangle$	0.5286(3)	0.0199(10)
$\langle W_{1 \times 2} \rangle$	0.3161(4)	0.0207(12)
$\langle W_{2 \times 4} \rangle$	0.0312(2)	0.0213(15)
$r_0/a$	3.87(17)	0.0302(10)
$C_\pi(0)$	2.132(2)	0.0230(18)
$C_\pi(1)$	0.2782(8)	0.026(3)
$C_\pi(2)$	0.0792(5)	0.030(5)
$C_\pi(3)$	0.0298(3)	0.035(80)
$m_\pi^{\text{eff}}(1)$	2.037(2)	0.030(5)
$m_\pi^{\text{eff}}(2)$	1.256(4)	0.035(28)

optimal matching for the lattice spacing (as opposed to the plaquette) would be  $\delta\beta=0.032(10)$ . This value is compared with those corresponding to the various Wilson loops in Fig. 1.

We conclude that the constant lattice volume curve (as defined by  $r_0$ ) may indeed differ from that corresponding to the constant plaquette value. With only 40 configurations, the evidence is of marginal statistical significance. As a cross-check we measured  $r_0$  directly from a simulation at (5.232, .1335). See Table V. The result was compatible with the prediction and with the matched ensemble from which the prediction was made. Again, the statistical significance is not high.

We have also measured  $r_0$  on quenched configurations at  $\beta=5.61$ , the point which is predicted and verified to have matched values of  $\langle P \rangle$ , and at  $\beta=5.85$ , close to the point where  $r_0$  is expected to match. Results are shown in Table III. As expected, the lattice spacing does *not* match at 5.61 but is close to matching at 5.85. The shift required to match  $r_0$  is some 60% larger than that required to match  $\langle P \rangle$ .

## F. Hadron correlators

Finally in this section, we present results from matching lattice pion correlators. This test is made more practicable by recent advances in measuring hadron correlators with good statistical precision on a single gauge configuration [13]. We have made measurements of the local pion correlator  $C_\pi(t)$  on the reference ensemble and calculated the corresponding  $\delta\beta$  shifts. That is, for each value of  $t$ , we estimate the shift  $\delta\beta$  required to keep  $C_\pi$  constant for the test  $\kappa$  shift. Results are presented Table VI where they are compared with examples of other observables. At short time separations (0 and 1) we find values compatible with that for the plaquette. At large values of  $t$  the results are overwhelmed by noise and we are unable to draw conclusions. However, there are indications that for increasing  $t$  the shift required in  $\beta$  is also increasing. See, for example, the correlator for  $t=2$  and corresponding effective mass values which are also shown in the table:

$$m_\pi^{\text{eff}}(t) = -\ln[C_\pi(t+1)/C_\pi(t)]. \quad (45)$$

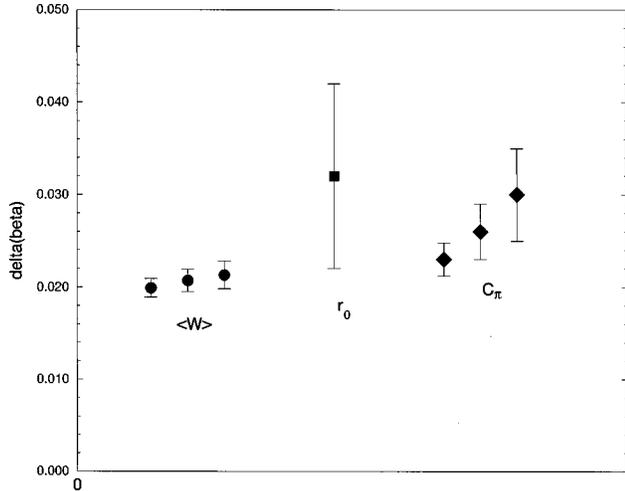


FIG. 4. Predictions for the shift  $\delta\beta$  from reference point (5.2, 0.1340) required to keep the specified observables constant when  $\kappa$  is changed by  $\delta\kappa = -0.0005$ . Circles correspond to Wilson loops ( $1 \times 1$ ,  $1 \times 2$ ,  $2 \times 4$ ), diamonds to  $C_\pi(t)$  ( $t=0,1,2$ ), and the square to  $r_0$ .

Clearly one requires greater statistics to confirm these trends, but they are consistent with those inferred from the  $r_0$  results presented above. A compilation of shifts  $\delta\beta$  for sample loops  $C_\pi(t)$  and  $r_0$  is shown in Fig. 4.

Note that the effective mass values involved in this test are very far from those of a physical pion.

## V. FURTHER APPLICATIONS

The above matching technology has a variety of potential uses including the following.

### A. Parallel tempering

Parallel tempering (PT) is an improved Monte Carlo method originally proposed by Hukushima and Nemoto [14] to improve simulations of spin glasses. It was further discussed by Marinari *et al.* in [15] and [16] who suggested its use in lattice QCD. Recently Boyd [17] applied the technique to lattice QCD with staggered fermions and found evidence that parallel tempering did help decorrelate long distance observables.

Parallel tempering essentially consists of running several independent simulations in parallel and with different parameters. Each such simulation produces a set of configurations which is distributed according to the probability distribution dictated by the simulation action and parameters. The PT algorithm exploits the fact that these distributions may have an overlap and occasionally attempts to swap configurations between ensembles. Acceptance of the swap is controlled by a Metropolis style acceptance step.

This is the same situation described by the matching criterion (M3) in Sec. II. From another viewpoint, the distance between the actions in matching criterion (M2) can be related to the acceptance rate of the swaps in a parallel tempering algorithm.

In [17], tempering was carried out in the quark mass only. All the ensembles had the same value of  $\beta$ . Furthermore, the quark masses had to be spaced quite closely together to obtain a reasonable swap acceptance rate. With the matching technology presented in this paper, it is possible to temper in *both*  $\beta$  and the quark mass. This may allow one to perform PT along a curve of approximately constant volume and at such a separation between ensembles that one might use some of the tempering ensembles to perform chiral extrapolations. Alternatively, one might be able to simulate with ensembles suitably chosen to improve the decorrelation properties of the system as a whole. A detailed investigation into PT using the matching technology is being conducted by us and the full results will be reported elsewhere [18].

### B. Approximate algorithms

In [5] we demonstrated how the parameters of approximate algorithms could be tuned according to the criteria (M1), (M2) or (M3) described above. In subsequent tests [7], we showed that approximate algorithms based on a few Wilson loops only were unlikely to produce a very accurate approximation, at least in the sense of (M3) where the approximate action acts as a guide within an exact algorithm. The variance of the difference between the two extensive quantities remains unacceptably large on lattices of a useful size. It is, however, still of considerable interest to design approximate or model actions which improve on the quenched approximation by encapsulating at least some of the additional physics implied by dynamical quarks.

An alternative route to an exact algorithm might be to make use of the Lanczos quadrature approximation of Sec. III for part of the effective action and gauge invariant loops as above for the remainder. The trial configurations would be generated by the loop part of the action and an accept-reject step based on the Lanczos part. The technology of Sec. II can be used to tune the loop part together with the Lanczos part to match the exact action. As a simple example of this approach, consider an approximate action defined such that [cf. Eq. (7)]

$$S_{\text{app}} = -\beta' W_{\square} - T(N_{\text{Lanc}}) \quad (46)$$

where  $T(N_{\text{Lanc}})$  is the approximation to  $\text{Tr} \ln(M^\dagger M)$  as described in Sec. III but using only  $N_{\text{Lanc}}$  Lanczos iterations. The loop part of the effective action is just a single plaquette in this example. In a standard Metropolis update this action would only be viable if  $N_{\text{Lanc}}$  was considerably smaller than the typical values (90) required to estimate the true action, and sufficient account was taken of the short range fluctuations by having a properly tuned gauge loop part. This approximation is similar in spirit to that advocated in [19,20] where it is argued that a truncated sum of low-lying eigenvalues can reproduce the gross behavior of the fermion determinant. In the present scheme, we are able to obtain a particularly efficient approximation to the trace logarithm by using the optimal weighting determined by the Gaussian quadrature rule. We have conducted preliminary tests of these ideas by measuring the shift  $\delta\beta = \beta' - \beta$  required to compensate for a truncation to  $N_{\text{Lanc}}$  Lanczos iterations. In a

simple test which matched the average plaquette, the shift in  $\beta$  was reduced by a factor of 8 in changing from  $N_{\text{Lanc}}=0$  (quenched) to  $N_{\text{Lanc}}=4$ , for example. The corresponding residual variance (after matching) also dropped by a factor of around 8. For increasing  $N_{\text{Lanc}}$ , the shift rapidly becomes compatible with zero at the level of statistical accuracy implied by the number of noise vectors used (80). This demonstrates that the Lanczos quadrature approach gives a very efficient estimator for the trace logarithm. It will be worth exploring whether the long range modes described by such an approximation can be combined with a suitably tuned gauge loop action describing short range modes so as to achieve a practical exact algorithm.

## VI. CONCLUSIONS

We have proposed a strategy for dynamical quark simulations in which the effective lattice volume is held fixed while the effects of progressively lighter sea quarks are investigated. Possible algorithms for accomplishing this have

been presented and the results of tests discussed. In particular, we have presented results using an efficient stochastic estimator of the fermion determinant and quantities related to it. These include estimates of the constant lattice spacing curves at relevant points in the  $\beta, \kappa$  plane for lattice QCD using a non-perturbatively improved action. We have demonstrated that the work involved in determining such curves via our differential stochastic methods is considerably less than that required to establish them by direct simulation. Further applications of these techniques have been discussed.

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